Recitation 2, Probability Review Paolo Bertolotti

1 Elementary Probability

Probability allows us to study the likelihood of an event's occurrence. Consider rolling a fair die. We do not know what number will be rolled with certainty but we know there is a chance of rolling a 4. Before assigning likelihoods to values, we define the basic underlying framework.

- Experiment: an action where the result is uncertain
- Sample space: the set of all possible outcomes of an experiment
- Event: a subset of the sample space

In our die example, the experiment is rolling a die once. The sample space is $\{1, 2, 3, 4, 5, 6\}$. Rolling a 4 would be an example of an event.

Given a sample space S, the **probability** P is a function from the space of events in S to the interval [0, 1]. It satisfies the following properties:

- 1. Countable additivity: For any sequence A_i of events in S such that $A_i \cap A_j = \emptyset$ for all $i \neq j$, then $P(\bigcup_{i=1}^{\infty} A_i) = \sum_{i=1}^{\infty} P(A_i)$. In words, the probability of any disjoint event occurring in some sequence is equal to the sum of their individual probabilities.
- 2. Normalization: P(S) = 1

Question: If $A \in S$ with probability P(A), what is the probability of the event A_c , A's complement? By their definition, A and A_c satisfy $A \cup A_c = S$ and $A \cap A_c = \emptyset$. Therefore, we have $1 = P(S) = P(A \cup A_c) = P(A) + P(A_c)$ and $P(A_c) = 1 - P(A)$. For example, the probability of not rolling a 4 is equal to 1 minus the probability of rolling a 4.

Given events A and B where P(B) > 0, the **conditional probability** of A given B is

$$P(A|B) = \frac{P(A \cap B)}{P(B)} \tag{1}$$

Example: Consider a fair die, where fair means every outcome in $\{1, 2, 3, 4, 5, 6\}$ happens with equal probability 1/6. Define the event $A = \{4or5\}$. What is P(A)? Simple, P(A) = 1/6 + 1/6 = 1/3. Now, what happens if we have additional information regarding the toss.

- Case 1: If $B_1 =$ "the outcome is an even number", then $P(A|B_1) = 1/3$
- Case 2: If $B_2 =$ "the outcome is larger than 3", then $P(A|B_2) = 2/3$

• Case 3: If $B_3 =$ "the outcome is less or equal to 3", then $P(A|B_3) = 0$

Two events are **independent** iff $P(A \cap B) = P(A)P(B)$. Independence implies P(A) = P(A|B) assuming P(B) > 0, which means that B's occurrence provides no information about A. Independence simplifies many calculations. Two important theorems using the concept of conditional expectation are:

Theorem 1 (Law of Total Probability) If A is an event and B_i is a sequence of n events that partitions the sample space (meaning they are all disjoint and their union equals the sample space), then

$$P(A) = \sum_{i=1}^{n} P(A|B_i) P(B_i).$$
 (2)

Theorem 2 (Bayes' Theorem) For events A and B with P(B) > 0,

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)} = \frac{P(B|A)P(A)}{P(B|A)P(A) + P(B|A_c)P(A_c)}$$
(3)

where the second equality follows from the Law of Total Probability.

2 Random Variables

A random variable (rv) is a function $X : S \to \mathbb{R}$ that assigns a real number to each event in the sample space. For example, consider an experiment where we toss a coin ten times. The sample space is the collection of all possible combinations of H and T of size 10. A possible event is $s = \{HHHTHHHTTT\} \in S$. We can define the random variable Y that counts the number of heads. In our example, Y(s) = 6.

A random variable is called **discrete** if it takes on at most a countable set of values. For every discrete random variable X we define the **probability mass function** (pmf) of X by

$$p_X(x) = P(\{s \in S : X(s) = x\})$$
(4)

We usually omit the argument of the v X and simply write

$$p_X(x) = P(X = x) \tag{5}$$

Assume the random variable X takes values in the set $a_1, a_2, ..., a_n$ and the random variable Y takes values in the set $b_1, b_2, ..., b_m$. We say that X and Y are **independent random variables** if

$$P(\{X = a_i\} \cap \{Y = b_j\}) = P(X = a_i)P(Y = b_j)$$
(6)

for every i = 1, 2, ..., n and j = 1, 2, ..., m.

All random variables have a **cumulative distribution function** F, which is defined as

$$F_X(x) = P(X \le x) \tag{7}$$

Continuous random variables, which can take a uncountably infinite set of values, do not have a pmf. Instead, they have a **probability density function** (pdf) f, defined as

$$f_X(x) = \frac{d}{dx}F(x) \tag{8}$$

where

$$P(x \in A) = \int_{A} f_X(x) dx \tag{9}$$

2.1 Expectation and Variance

The **expected value** of a random variable X, also called its mean, is the probability-weighted average value for the variable. It is defined as

$$E[X_{continuous}] = \int x f_X(x) dx \tag{10}$$

$$E[X_{discrete}] = \sum_{i} x_i \, p_X(x_i) \tag{11}$$

For independent random variables X and Y, E[XY] = E[X]E[Y].

The **variance** of a random variable X measures how dispersed it is around its mean. In a sense, it captures how variable it is.

$$Var(X) = E[(X - E[X])^2] = E[X^2] - E[X]^2$$
(12)

The **covariance** of two random variables X and Y measures how much they co-vary or co-move. I.e., if X goes up, does Y go up or down. It is defined as

$$cov(X,Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y]$$
 (13)

Theorem 3 (Variance of Sums) For any sequence X_i of random variables

$$Var(\sum_{i=1}^{n} X_{i}) = \sum_{i=1}^{n} Var(X_{i}) + \sum_{i \neq j} cov(X_{i}, X_{j})$$
(14)

Remark: As a special case, when X_i and X_j are independent for each $i \neq j$, the second term above vanishes and

$$Var(\sum_{i=1}^{n} X_i) = \sum_{i=1}^{n} Var(X_i)$$
 (15)

2.2 Common Distributions

A quick reminder of some useful random variables and their distributions.

• Bernoulli: Models a biased coin toss with bias parameter p. Random variable X takes value 1 with probability p and value 0 with probability 1 - p

- **Binomial:** Models the sum of outcomes of *n* biased coin tosses with bias parameter *p*. Random variable X takes the values in $\{0, ..., n\}$ with pmf $p_X(x) = \binom{n}{x} p^x (1-p)^{n-x}$
- Geometric: Models the number of tosses until a heads appears in biased coin tosses with bias parameter p (giving the probability of heads). I.e. the number of trials until the first success. Random variable X takes the values in $\{1, ...\}$ with pmf $p_X(x) = (1-p)^{x-1}p$
- **Poisson:** One form of continuum limit for the binomial distribution when p becomes small and n becomes large. Governed by an intensity / rate parameter λ . Random variable X takes values in $\{0, ...\}$ with pmf $p_X(x) = \frac{e^{-\lambda}\lambda^x}{x!}$

¹The skeleton for these notes was provided by the recitation notes for MIT 6.268 Network Science and Models

1.022 Introduction to Network Models Fall 2018

For information about citing these materials or our Terms of Use, visit: <u>https://ocw.mit.edu/terms</u>.