

14.31/14.310 Lecture 9

Probability---moments of a distribution

Ok, back to probability now.

Where were we? Ah, yes, talking about moments of distributions, expectation, in particular.

Probability---moments of a distribution

What if, instead of wanting to know a certain feature of the distribution of X , say expectation, we are interested, instead in that feature of the distribution of $Y = g(X)$.

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Well, we can obviously figure out how Y is distributed--we know how to do that--and then use that distribution to compute, say, $E(Y)$.

There may be an easier way--it can be shown that

$$E(Y) = E(g(X)) = \int y f_Y(y) dy = \int g(x) f_X(x) dx$$

Probability---St. Petersburg paradox

Classic example/paradox in probability theory, but one where economists come out looking particularly good.

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Here's the game: I flip a fair coin until it comes up heads. If the number of flips necessary is X , I pay you 2^X dollars. How much would you be willing to pay me to play this game?

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So let's calculate them:

Let X = number of flips required.

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No one would be willing to pay me an infinite amount to play this game.

I would guess that I wouldn't have any takers at \$20, and that's a lot less than infinity.

That's the paradox, but is it really?

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Economists know that people have diminishing marginal utility of money. In other words, their valuation of additional money decreases as the amount of money they have increases.

So let Z = valuation of winnings = $\log(Y) = \log(2^X)$

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Then, $E(Z) = \sum_x \log(2^x)(1/2)^x$

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So let $Z =$ valuation of winnings $= \log(Y) = \log(2^X)$

$$\begin{aligned} \text{Then, } E(Z) &= \sum_x \log(2^x) (1/2)^x \\ &= \log(2) \sum_x x (1/2)^x \end{aligned}$$

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$$\begin{aligned} \text{Then, } E(Z) &= \sum_x \log(2^x) (1/2)^x \\ &= \log(2) \sum_x x (1/2)^x \\ &= 2 \log(2) < \infty \end{aligned}$$

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So this is only a paradox unless you know a little bit of economics.

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Really, what if the X 's aren't independent? Yes, really, they don't have to be independent.

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5. $E(XY) = E(X)E(Y)$ if X, Y independent

Probability---another moment: variance

In addition to describing the location, or center, of a distribution of a random variable, we often would like to describe how spread out it is. There's a moment for that, variance.

$$\text{Var}(X) = E[(X-\mu)^2]$$

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Note that variance is an expectation, so many of its properties will follow from that.

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In other words, shift a distribution and its variance doesn't change.
Shrink or spread out a distribution and its variance changes by the square of the multiplicative factor.

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Ah, here we actually need independence.

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6. $\text{Var}(X) = E(X^2) - [E(X)]^2$

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This last property can provide a handy way to compute variance.

Probability---standard deviation

Often it's convenient for the measure of dispersion to have the same units as the random variable. For this reason, we define standard deviation.

$$SD(X) = \sigma = \sqrt{\text{Var}(X)} = \sqrt{\sigma^2}$$

Probability---variance of a function

Since variance is an expectation, we can apply the results of expectation of a function of a random variable to get variance of a function of a random variable.

So if $Y = r(X)$,

$$\begin{aligned}\text{Var}(Y) &= E(Y^2) - E(Y)^2 = E(r(x)^2) - E(r(x))^2 \\ &= \int r(x)^2 f_x(x) dx - \left[\int r(x) f_x(x) dx \right]^2\end{aligned}$$

Probability---conditional expectation

A conditional expectation is the expectation of a conditional distribution. In other words,

$$E(Y|X) = \int y f_{Y|X}(y|x) dy$$

Note that $E(Y|X)$ is a function of X , and, therefore, a random variable. $E(Y|X=x)$ is just a number.

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Thm $E(E(Y|X)) = E(Y)$ "Law of Iterated Expectations"

Probability---conditional variance

The definition of conditional variance follows from that of variance and conditional expectation.

Thm $\text{Var}(E(Y|X)) + E(\text{Var}(Y|X)) = \text{Var}(Y)$

"Law of Total Variance"

Probability---two laws

"Law of Iterated Expectations"

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May seem a little mysterious, not clear how they're useful.

Probability---example

A former student of mine started an innovation incubator in NYC. Suppose he's been doing this for a few years and has kept track of the number of patents produced every year in his incubator. He knows that $E(N) = 2$ and $\text{Var}(N) = 2$.

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Suppose there are 5 patents this year. What is the probability that 3 are commercial successes?

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$$S|N=n \sim B(n, .2),$$

$$\text{so } P(S=3|N=5) = 5! / (3!2!) \cdot 2^3(1-.2)^2$$

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How do we get this? Compute the expectation of a Bernoulli random variable and add it up n times

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$$E(S) = E(E(S|N)) = E(Np) = .2E(N) = .4$$

Probability---example

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What is the (unconditional) variance of number of commercial successes? Can use the Law of Total Variance.

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Probability---example

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$$\begin{aligned}\text{Var}(S) &= \text{Var}(E(S|N)) + E(\text{Var}(S|N)) \\ &= \text{Var}(Np) + E(Np(1-p)) \\ &= .2^2\text{Var}(N) + .2(1-.2)E(N) = .4\end{aligned}$$

Probability---covariance and correlation

We now have moments to describe the location, or center, of a distribution of a random variable and how spread out that distribution is. We are often interested in the relationship between random variables, and we have a moment of joint distributions to describe one aspect of that relationship, covariance.

$$\text{Cov}(X, Y) = E[(X - \mu_X)(Y - \mu_Y)]$$

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We often denote $\text{Cov}(X, Y)$ with σ_{XY}

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$$\text{Cov}(X, Y) = E[(X - \mu_X)(Y - \mu_Y)]$$

And we have a standardized version, correlation.

$$\rho(X, Y) = E[(X - \mu_X)(Y - \mu_Y)] / \sqrt{\text{Var}(X)} \sqrt{\text{Var}(Y)}$$

Probability---covariance and correlation

$$\rho(X, Y) = E[(X - \mu_X)(Y - \mu_Y)] / \sqrt{\text{Var}(X)} \sqrt{\text{Var}(Y)}$$

We say that $X \xi Y$ are "positively correlated" if $\rho > 0$.

We say that $X \xi Y$ are "negatively correlated" if $\rho < 0$.

We say that $X \xi Y$ are "uncorrelated" if $\rho = 0$.

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7. $|\rho(X, Y)| \leq 1$

8. $|\rho(X, Y)| = 1$ iff $Y = aX + b, a \neq 0$

Probability---a preview of regression

We have two random variables, X & Y .

$$EX = \mu_X, \text{Var}X = \sigma_X^2$$

$$EY = \mu_Y, \text{Var}Y = \sigma_Y^2$$

$$\rho_{XY} = \text{Cov}(X, Y) / (\sigma_X \sigma_Y)$$

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$$\rho_{XY} = \text{Cov}(X, Y) / (\sigma_X \sigma_Y)$$

We know that, if $\rho_{XY} = 1$ then $Y = a + bX$, $b > 0$, and if $\rho_{XY} = -1$ then $Y = a + bX$, $b < 0$.

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$$EY = \mu_Y, \text{Var}Y = \sigma_Y^2$$

$$\rho_{XY} = \text{Cov}(X, Y) / (\sigma_X \sigma_Y)$$

We know that, if $\rho_{XY} = 1$ then $Y = a + bX$, $b > 0$, and if

$\rho_{XY} = -1$ then $Y = a + bX$, $b < 0$.

If $|\rho_{XY}| < 1$, then we can write $Y = \alpha + \beta X + V$.

Probability---a preview of regression

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V is another random variable,
but what can we say about it?

Probability---a preview of regression

What we can say about V depends on how we define α & β .

$$\text{Let } \beta = \rho_{XY}\sigma_Y/\sigma_X$$

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$E(V) = 0$ and $\text{Cov}(X, V) = 0$. (You can show this easily using properties of expectation, variance, and covariance that we've seen.)

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We then call α & β "regression coefficients," and think of $\alpha + \beta X$ as the part of Y "explained by" X and V as the "unexplained" part.

Probability---inequalities

Two inequalities involving moments of distributions and tail probabilities often come in handy:

Markov Inequality

X is a random variable that is always non-negative.

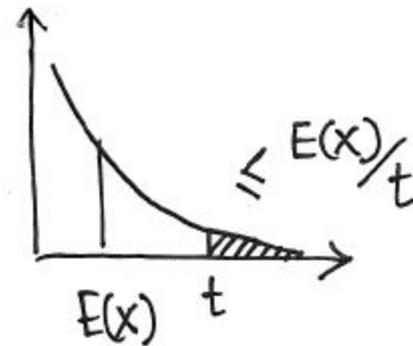
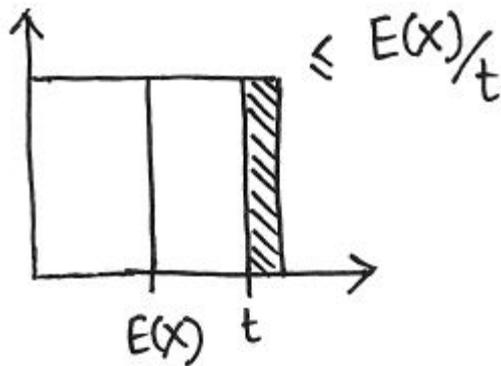
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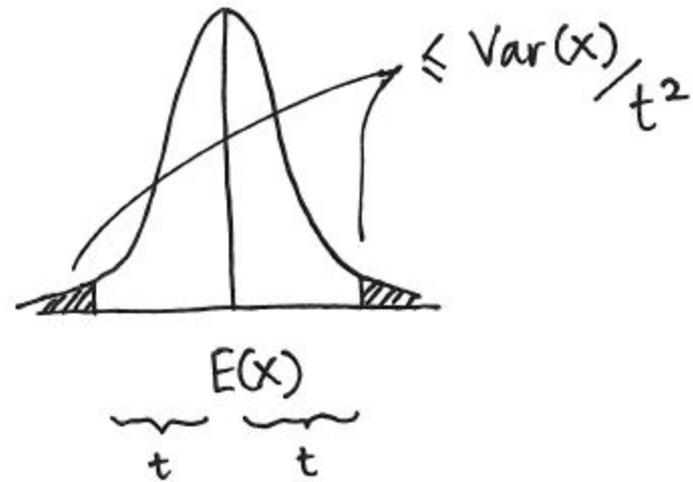
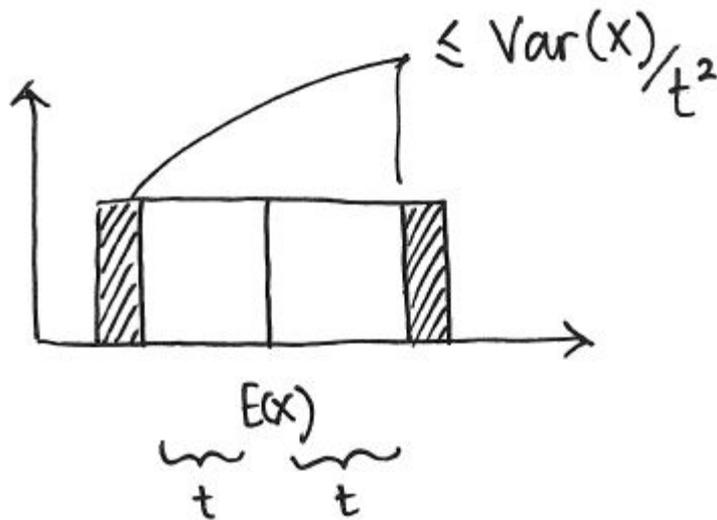
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X is a random variable for which $\text{Var}(X)$ exists. Then for any $t > 0$, $P(|X - E(X)| \geq t) \leq \text{Var}(X)/t^2$.

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