

## 3 Alterations

### 3.1 Ramsey numbers

Recall from [Section 1.1](#):

$R(s, t)$  = smallest  $n$  such that every red/blue edge coloring of  $K_n$  contains a red  $K_s$  or a blue  $K_t$

Using the basic method (union bounds), we deduce

**Theorem 3.1.1.** If there exists  $p \in [0, 1]$  with

$$\binom{n}{s} p^{\binom{s}{2}} + \binom{n}{t} (1-p)^{\binom{t}{2}} < 1$$

then  $R(s, t) > n$ .

*Proof sketch.* Color edge red with prob  $p$  and blue with prob  $1-p$ . LHS upper bounds the probability of a red  $K_s$  or a blue  $K_t$ .  $\square$

Using the alteration method, we deduce

**Theorem 3.1.2.** For all  $p \in [0, 1]$  and  $n$ ,

$$R(s, t) > n - \binom{n}{s} p^{\binom{s}{2}} - \binom{n}{t} (1-p)^{\binom{t}{2}}$$

*Proof sketch.* Color edge red with prob  $p$  and blue with prob  $1-p$  remove one vertex from each red  $K_s$  or blue  $K_t$ . RHS lower bounds the expected number remaining vertices.  $\square$

### 3.2 Dominating set in graphs

In a graph  $G = (V, E)$ , we say that  $U \subset V$  is **dominating** if every vertex in  $V \setminus U$  has a neighbor in  $U$ .

**Theorem 3.2.1.** Every graph on  $n$  vertices with minimum degree  $\delta > 1$  has a dominating set of size at most  $\left(\frac{\log(\delta+1)+1}{\delta+1}\right)n$ .

Naive attempt: take out vertices greedily. The first vertex eliminates  $1 + \delta$  vertices, but subsequent vertices eliminate possibly fewer vertices.

*Proof.* Two-step process (alteration method):

1. Choose a random subset
2. Add enough vertices to make it dominating

Let  $p \in [0, 1]$  to be decided later. Let  $X$  be a random subset of  $V$  where every vertex is included with probability  $p$  independently.

Let  $Y = V \setminus (X \cup N(X))$ . Each  $v \in V$  lies in  $Y$  with probability  $\leq (1 - p)^{1+\delta}$ .

Then  $X \cup Y$  is dominating, and

$$\mathbb{E}[|X \cup Y|] = \mathbb{E}[|X|] + \mathbb{E}[|Y|] \leq pn + (1 - p)^{1+\delta}n \leq (p + e^{-p(1+\delta)})n$$

using  $1 + x \leq e^x$  for all  $x \in \mathbb{R}$ . Finally, setting  $p = \frac{\log(\delta+1)}{\delta+1}$  to minimize  $p + e^{-p(1+\delta)}$ , we bound the above expression by

$$\leq \left( \frac{1 + \log(\delta + 1)}{\delta + 1} \right). \quad \square$$

### 3.3 Heilbronn triangle problem

**Question 3.3.1.** How can one place  $n$  points in the unit square so that no three points forms a triangle with small area?

Let

$$\Delta(n) = \sup_{\substack{S \subset [0,1]^2 \\ |S|=n}} \min_{\substack{p,q,r \in S \\ \text{distinct}}} \text{area}(pqr)$$

Naive constructions fair poorly. E.g.,  $n$  points around a circle has a triangle of area  $\Theta(1/n^3)$  (the triangle formed by three consecutive points has side lengths  $\asymp 1/n$  and angle  $\theta = (1 - 1/n)2\pi$ ). Even worse is arranging points on a grid, as you would get triangles of zero area.

Heilbronn conjectured that  $\Delta(n) = O(n^{-2})$ .

Komlós, Pintz, and Szemerédi (1982) disproved the conjecture, showing  $\Delta(n) \gtrsim n^{-2} \log n$ . They used an elaborate probabilistic construction. Here we show a much simpler version probabilistic construction that gives a weaker bound  $\Delta(n) \gtrsim n^{-2}$ .

*Remark 3.3.2.* The currently best upper bound known is  $\Delta(n) \leq n^{-8/7+o(1)}$  (Komlós, Pintz, and Szemerédi 1981)

**Theorem 3.3.3.** For every positive integer  $n$ , there exists a set of  $n$  points in  $[0, 1]^2$  such that every triple spans a triangle of area  $\geq cn^{-2}$ , for some absolute constant  $c > 0$ .

*Proof.* Choose  $2n$  points at random. For every three random points  $p, q, r$ , let us estimate

$$\mathbb{P}_{p,q,r}(\text{area}(p, q, r) \leq \epsilon).$$

By considering the area of a circular annulus around  $p$ , with inner and outer radii  $x$  and  $x + \Delta x$ , we find



$$\mathbb{P}_{p,q}(|pq| \in [x, x + \Delta x]) \leq \pi((x + \Delta x)^2 - x^2)$$

So the probability density function satisfies

$$\mathbb{P}_{p,q}(|pq| \in [x, x + dx]) \leq 2\pi x dx$$

For fixed  $p, q$

$$\mathbb{P}_r(\text{area}(pqr) \leq \epsilon) = \mathbb{P}_r\left(\text{dist}(pq, r) \leq \frac{2\epsilon}{|pq|}\right) \lesssim \frac{\epsilon}{|pq|}$$

Thus, with  $p, q, r$  at random

$$\mathbb{P}_{p,q,r}(\text{area}(pqr) \leq \epsilon) \lesssim \int_0^{\sqrt{2}} 2\pi x \frac{\epsilon}{x} dx \asymp \epsilon.$$

Given these  $2n$  random points, let  $X$  be the number of triangles with area  $\leq \epsilon$ . Then  $\mathbb{E}X = O(\epsilon n^3)$ .

Choose  $\epsilon = c/n^2$  with  $c > 0$  small enough so that  $\mathbb{E}X \leq n$ .

Delete a point from each triangle with area  $\leq \epsilon$ .

The expected number of remaining points is  $\mathbb{E}[2n - X] \geq n$ , and no triangles with area  $\leq \epsilon = c/n^2$ .

Thus with positive probability, we end up with  $\geq n$  points and no triangle with area  $\leq c/n^2$ .  $\square$

**Algebraic construction.** Here is another construction due to Erdős (in appendix of Roth (1951)) also giving  $\Delta(n) \gtrsim n^{-2}$ :

Let  $p$  be a prime. The set  $\{(x, x^2) \in \mathbb{F}_p^2 : x \in \mathbb{F}_p\}$  has no 3 points collinear (a parabola meets every line in  $\leq 2$  points). Take the corresponding set of  $p$  points in  $[p]^2 \subset \mathbb{Z}^2$ . Then every triangle has area  $\geq 1/2$  due to Pick's theorem. Scale back down to a unit square. (If  $n$  is not a prime, then use that there is a prime between  $n$  and  $2n$ .)

### 3.4 Markov's inequality

We note an important tool that will be used next.

**Markov's inequality.** Let  $X \geq 0$  be random variable. Then for every  $a > 0$ ,

$$\mathbb{P}(X \geq a) \leq \frac{\mathbb{E}[X]}{a}.$$

*Proof.*  $\mathbb{E}[X] \geq \mathbb{E}[X1_{X \geq a}] \geq \mathbb{E}[a1_{X \geq a}] = a\mathbb{P}(X \geq a)$  □

Take-home message: for r.v.  $X \geq 0$ , if  $\mathbb{E}X$  is *very* small, then *typically*  $X$  is small.

### 3.5 High girth and high chromatic number

If a graph has a  $k$ -clique, then you know that its chromatic number is at least  $k$ .

Conversely, if a graph has high chromatic number, is it always possible to certify this fact from some “local information”?

Surprisingly, the answer is no. The following ingenious construction shows that a graph can be “locally tree-like” while still having high chromatic number.

The **girth** of a graph is the length of its shortest cycle.

**Theorem 3.5.1 (Erdős 1959).** For all  $k, \ell$ , there exists a graph with girth  $> \ell$  and chromatic number  $> k$ .

*Proof.* Let  $G \sim G(n, p)$  with  $p = (\log n)^2/n$  (the proof works whenever  $\log n/n \ll p \ll n^{-1+1/\ell}$ ). Here  $G(n, p)$  is Erdős–Rényi random graph ( $n$  vertices, every edge appearing with probability  $p$  independently).

Let  $X$  be the number of cycles of length at most  $\ell$  in  $G$ . By linearity of expectations, as there are exactly  $\binom{n}{i}(i-1)!/2$  cycles of length  $i$  in  $K_n$  for each  $3 \leq i \leq n$ , we have (recall that  $\ell$  is a constant)

$$\mathbb{E}X = \sum_{i=3}^{\ell} \binom{n}{i} \frac{(i-1)!}{2} p^i \leq \sum_{i=3}^{\ell} n^i p^i = o(n).$$

By Markov's inequality

$$\mathbb{P}(X \geq n/2) \leq \frac{\mathbb{E}X}{n/2} = o(1).$$

(This allows us to get rid of all short cycles.)

How can we lower bound the chromatic number  $\chi(\cdot)$ ? Note that  $\chi(G) \geq |V(G)|/\alpha(G)$ , where  $\alpha(G)$  is the independence number (the size of the largest independent set).

With  $x = (3/p) \log n$ ,

$$\mathbb{P}(\alpha(G) \geq x) \leq \binom{n}{x} (1-p)^{\binom{x}{2}} < n^x e^{-px(x-1)/2} = (ne^{-p(x-1)/2})^x = o(1).$$

Let  $n$  be large enough so that  $\mathbb{P}(X \geq n/2) < 1/2$  and  $\mathbb{P}(\alpha(G) \geq x) < 1/2$ . Then there is some  $G$  with fewer than  $n/2$  cycles of length  $\leq \ell$  and with  $\alpha(G) \leq (3/p) \log n$ .

Remove a vertex from each cycle to get  $G'$ . Then  $|V(G')| \geq n/2$ , girth  $> \ell$ , and  $\alpha(G') \leq \alpha(G) \leq (3/p) \log n$ , so

$$\chi(G') \geq \frac{|V(G')|}{\alpha(G')} \geq \frac{np}{6 \log n} = \frac{\log n}{6} > k$$

if  $n$  is sufficiently large. □

*Remark 3.5.2.* Erdős (1962) also showed that in fact one needs to see at least a linear number of vertices to deduce high chromatic number: for all  $k$ , there exists  $\epsilon = \epsilon_k$  such that for all sufficiently large  $n$  there exists an  $n$ -vertex graph with chromatic number  $> k$  but every subgraph on  $\lfloor \epsilon n \rfloor$  vertices is 3-colorable. (In fact, one can take  $G \sim G(n, C/n)$ ; see "Probabilistic Lens: Local coloring" in Alon–Spencer)

### 3.6 Greedy random coloring

Recall  $m(k)$  is the minimum number of edges in a  $k$ -uniform hypergraph that is not 2-colorable.

Earlier we proved that  $m(k) \geq 2^{k-1}$ . Indeed, given a  $k$ -graph with  $< 2^{k-1}$  edges, by randomly coloring the vertices, the expected number of monochromatic numbers is  $< 1$ .

We also proved an upper bound  $m(k) = O(k^2 2^k)$  by taking a random  $k$ -uniform hypergraph on  $k^2$  vertices.

Here is the currently best known lower bound.

**Theorem 3.6.1** (Radhakrishnan and Srinivasan (2000)).  $m(k) \gtrsim \sqrt{\frac{k}{\log k}} 2^k$

Here we present a simpler proof, based on a **random greedy coloring**, due to Cherkashin and Kozik (2015), following an approach of Pluhaár (2009).

*Proof.* Suppose  $H$  is a  $k$ -graph with  $m$  edges.

Map  $V(H) \rightarrow [0, 1]$  uniformly at random.

Color vertices greedily from left to right: color a vertex blue unless it would create a monochromatic edge, in which case color it red (i.e., every red vertex is the final vertex in an edge with all earlier  $k - 1$  vertices have been colored blue).

The resulting coloring has no all-blue edges. What is the probability of seeing a red edge?

If there is a red edge, then there must be two edges  $e, f$  so that the last vertex of  $e$  is the first vertex of  $f$ . Call such pair  $(e, f)$  **conflicting**.

Want to bound probability of seeing a conflicting pair in a random  $V(H) \rightarrow [0, 1]$ .

Here is an attempt (an earlier weaker result due to [Pluhaár \(2009\)](#)). Each pair of edges with exactly one vertex in common conflicts with probability  $\frac{(k-1)!^2}{(2k-1)!} = \frac{1}{2k-1} \binom{2k-2}{k-1}^{-1} \asymp k^{-1/2} 2^{-2k}$ ; union bounding over  $< m^2$  pairs of edges, the probability of getting a conflicting edge is  $\lesssim m^2 k^{-1/2} 2^{-2k}$ , which is  $< 1$  for some  $m \asymp k^{1/4} 2^k$ .

We'd like to do better by more carefully analyzing conflicting edges. Continuing ...

Write  $[0, 1] = L \cup M \cup R$  where ( $p$  to be decided)

$$L := \left[0, \frac{1-p}{2}\right) \quad M := \left[\frac{1-p}{2}, \frac{1+p}{2}\right] \quad R := \left(\frac{1+p}{2}, 1\right].$$

The probability that a given edge lands entirely in  $L$  is  $(\frac{1-p}{2})^k$ , and likewise with  $R$

So probability that some edge of  $H$  is entirely contained in  $L$  or contained in  $R$  is  $\leq 2m(\frac{1-p}{2})^k$ .

Suppose that no edge of  $H$  lies entirely in  $L$  or entirely in  $R$ . If  $(e, f)$  conflicts, then their unique common vertex  $x_v \in e \cap f$  must lie in  $M$ . So the probability that  $(e, f)$  conflicts is (here we use  $x(1-x) \leq 1/4$ )

$$\int_{(1-p)/2}^{(1+p)/2} x^{k-1} (1-x)^{k-1} dx \leq p 4^{-k+1}.$$

Thus the probability of seeing any conflicting pair is

$$\leq 2m \left(\frac{1-p}{2}\right)^k + m^2 p 4^{-k+1} < 2^{-k+1} m e^{-pk} + (2^{-k+1} m)^2 p.$$

Set  $p = \log(2^{-k+2} k/m)/k$ , we find that the above probability is  $< 1$  for  $m = c 2^k \sqrt{k/\log k}$ , with  $c > 0$  being a sufficiently small constant.  $\square$

MIT OpenCourseWare  
<https://ocw.mit.edu>

18.226 Probabilistic Method in Combinatorics  
Fall 2020

For information about citing these materials or our Terms of Use, visit: <https://ocw.mit.edu/terms>.