

For a fixed  $f \in \mathcal{F}$ , if we observe  $\frac{1}{n} \sum_{i=1}^n I(f(X_i) \neq Y_i)$  is small, can we say that  $\mathbb{P}(f(X) \neq Y)$  is small? By the Law of Large Numbers,

$$\frac{1}{n} \sum_{i=1}^n I(f(X_i) \neq Y_i) \rightarrow \mathbb{E}I(f(X) \neq Y) = \mathbb{P}(f(X) \neq Y).$$

The Central Limit Theorem says

$$\frac{\sqrt{n} \left( \frac{1}{n} \sum_{i=1}^n I(f(X_i) \neq Y_i) - \mathbb{E}I(f(X) \neq Y) \right)}{\sqrt{\text{Var}I}} \rightarrow \mathcal{N}(0, 1).$$

Thus,

$$\frac{1}{n} \sum_{i=1}^n I(f(X_i) \neq Y_i) - \mathbb{E}I(f(X) \neq Y) \sim \frac{k}{\sqrt{n}}.$$

Let  $Z_1, \dots, Z_n \in \mathbb{R}$  be i.i.d. random variables. We're interested in bounds on  $\frac{1}{n} \sum Z_i - \mathbb{E}Z$ .

- (1) Jensen's inequality: If  $\phi$  is a convex function, then  $\phi(\mathbb{E}Z) \leq \mathbb{E}\phi(X)$ .
- (2) Chebyshev's inequality: If  $Z \geq 0$ , then  $\mathbb{P}(Z \geq t) \leq \frac{\mathbb{E}Z}{t}$ .

Proof:

$$\begin{aligned} \mathbb{E}Z &= \mathbb{E}ZI(Z < t) + \mathbb{E}ZI(Z \geq t) \geq \mathbb{E}ZI(Z \geq t) \\ &\geq \mathbb{E}tI(Z \geq t) = t\mathbb{P}(Z \geq t). \end{aligned}$$

- (3) Markov's inequality: Let  $Z$  be a signed r.v. Then for any  $\lambda > 0$

$$\mathbb{P}(Z \geq t) = \mathbb{P}(e^{\lambda Z} \geq e^{\lambda t}) \leq \frac{\mathbb{E}e^{\lambda Z}}{e^{\lambda t}}$$

and therefore

$$\mathbb{P}(Z \geq t) \leq \inf_{\lambda > 0} e^{-\lambda t} \mathbb{E}e^{\lambda Z}.$$

**Theorem 5.1** (Bennett). Assume  $\mathbb{E}Z = 0$ ,  $\mathbb{E}Z^2 = \sigma^2$ ,  $|Z| < M = \text{const}$ ,  $Z_1, \dots, Z_n$  independent copies of  $Z$ , and  $t \geq 0$ . Then

$$\mathbb{P}\left(\sum_{i=1}^n Z_i \geq t\right) \leq \exp\left(-\frac{n\sigma^2}{M^2} \phi\left(\frac{tM}{n\sigma^2}\right)\right),$$

where  $\phi(x) = (1+x)\log(1+x) - x$ .

*Proof.* Since  $Z_i$  are i.i.d.,

$$\mathbb{P}\left(\sum_{i=1}^n Z_i \geq t\right) \leq e^{-\lambda t} \mathbb{E} e^{\lambda \sum_{i=1}^n Z_i} = e^{-\lambda t} \prod_{i=1}^n \mathbb{E} e^{\lambda Z_i} = e^{-\lambda t} (\mathbb{E} e^{\lambda Z})^n.$$

Expanding,

$$\begin{aligned} \mathbb{E} e^{\lambda Z} &= \mathbb{E} \sum_{k=0}^{\infty} \frac{(\lambda Z)^k}{k!} = \sum_{k=0}^{\infty} \lambda^k \frac{\mathbb{E} Z^k}{k!} \\ &= 1 + \sum_{k=2}^{\infty} \frac{\lambda^k}{k!} \mathbb{E} Z^2 Z^{k-2} \leq 1 + \sum_{k=2}^{\infty} \frac{\lambda^k}{k!} M^{k-2} \sigma^2 \\ &= 1 + \frac{\sigma^2}{M^2} \sum_{k=2}^{\infty} \frac{\lambda^k M^k}{k!} = 1 + \frac{\sigma^2}{M^2} (e^{\lambda M} - 1 - \lambda M) \\ &\leq \exp\left(\frac{\sigma^2}{M^2} (e^{\lambda M} - 1 - \lambda M)\right) \end{aligned}$$

where the last inequality follows because  $1 + x \leq e^x$ .

Combining the results,

$$\begin{aligned} \mathbb{P}\left(\sum_{i=1}^n Z_i \geq t\right) &\leq e^{-\lambda t} \exp\left(\frac{n\sigma^2}{M^2} (e^{\lambda M} - 1 - \lambda M)\right) \\ &= \exp\left(-\lambda t + \frac{n\sigma^2}{M^2} (e^{\lambda M} - 1 - \lambda M)\right) \end{aligned}$$

Now, minimize the above bound with respect to  $\lambda$ . Taking derivative w.r.t.  $\lambda$  and setting it to zero:

$$-t + \frac{n\sigma^2}{M^2} (Me^{\lambda M} - M) = 0$$

$$e^{\lambda M} = \frac{tM}{n\sigma^2} + 1$$

$$\lambda = \frac{1}{M} \log\left(1 + \frac{tM}{n\sigma^2}\right).$$

The bound becomes

$$\begin{aligned}
 \mathbb{P}\left(\sum_{i=1}^n Z_i \geq t\right) &\leq \exp\left(-\frac{t}{M} \log\left(1 + \frac{tM}{n\sigma^2}\right) + \frac{n\sigma^2}{M^2} \left(\frac{tM}{n\sigma^2} + 1 - \log\left(1 + \frac{tM}{n\sigma^2}\right)\right)\right) \\
 &= \exp\left(\frac{n\sigma^2}{M^2} \left(\frac{tM}{n\sigma^2} - \log\left(1 + \frac{tM}{n\sigma^2}\right) - \frac{tM}{n\sigma^2} \log\left(1 + \frac{tM}{n\sigma^2}\right)\right)\right) \\
 &= \exp\left(\frac{n\sigma^2}{M^2} \left(\frac{tM}{n\sigma^2} - \left(1 + \frac{tM}{n\sigma^2}\right) \log\left(1 + \frac{tM}{n\sigma^2}\right)\right)\right) \\
 &= \exp\left(-\frac{n\sigma^2}{M^2} \phi\left(\frac{tM}{n\sigma^2}\right)\right)
 \end{aligned}$$

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