18.600: Lecture 15 Lectures 1-14 Review

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Counting tricks and basic principles of probability

Discrete random variables

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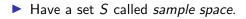
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- How many sequences a₁,..., a_k of non-negative integers satisfy a₁ + a₂ + ... + a_k = n?
- Answer: ^{n+k-1} n. Represent partition by k - 1 bars and n stars, e.g., as ** | ** || ** ** |*.



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- Countable additivity: $P(\bigcup_{i=1}^{\infty} E_i) = \sum_{i=1}^{\infty} P(E_i)$ if $E_i \cap E_j = \emptyset$ for each pair *i* and *j*.

▶
$$P(A^c) = 1 - P(A)$$

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•
$$A \subset B$$
 implies $P(A) \leq P(B)$

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More generally,

$$P(\bigcup_{i=1}^{n} E_{i}) = \sum_{i=1}^{n} P(E_{i}) - \sum_{i_{1} < i_{2}} P(E_{i_{1}}E_{i_{2}}) + \dots + (-1)^{(r+1)} \sum_{i_{1} < i_{2} < \dots < i_{r}} P(E_{i_{1}}E_{i_{2}} \dots E_{i_{r}})$$
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► The notation ∑_{i1 < i2 < ... < ir} means a sum over all of the ⁿ_r subsets of size r of the set ¹₂₂, 2, ..., n}.

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- $P(\cup_{i=1}^{n} E_i) = 1 \frac{1}{2!} + \frac{1}{3!} \frac{1}{4!} + \dots \pm \frac{1}{n!}$
- ► $1 P(\bigcup_{i=1}^{n} E_i) = 1 1 + \frac{1}{2!} \frac{1}{3!} + \frac{1}{4!} \dots \pm \frac{1}{n!} \approx 1/e \approx .36788$

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- Nice fact: $P(E_1E_2E_3...E_n) = P(E_1)P(E_2|E_1)P(E_3|E_1E_2)...P(E_n|E_1...E_{n-1})$

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- Useful when we think about multi-step experiments.
- For example, let E_i be event ith person gets own hat in the n-hat shuffle problem.



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In words: want to know the probability of *E*. There are two scenarios *F* and *F^c*. If I know the probabilities of the two scenarios and the probability of *E* conditioned on each scenario, I can work out the probability of *E*.

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▶ Ratio
$$\frac{P(B|A)}{P(B)}$$
 determines "how compelling new evidence is".

We can check the probability axioms: 0 ≤ P(E|F) ≤ 1, P(S|F) = 1, and P(∪E_i) = ∑ P(E_i|F), if *i* ranges over a countable set and the E_i are disjoint. ▶ We can check the probability axioms: $0 \le P(E|F) \le 1$, P(S|F) = 1, and $P(\cup E_i) = \sum P(E_i|F)$, if *i* ranges over a countable set and the E_i are disjoint.

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- The probability measure $P(\cdot|F)$ is related to $P(\cdot)$.
- To get former from latter, we set probabilities of elements outside of F to zero and multiply probabilities of events inside of F by 1/P(F).
- ▶ P(·) is the prior probability measure and P(·|F) is the posterior measure (revised after discovering that F occurs).

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- Say *E* and *F* are **independent** if P(EF) = P(E)P(F).
- Equivalent statement: P(E|F) = P(E). Also equivalent: P(F|E) = P(F).

Say $E_1 \dots E_n$ are independent if for each $\{i_1, i_2, \dots, i_k\} \subset \{1, 2, \dots n\}$ we have $P(E_{i_1}E_{i_2}\dots E_{i_k}) = P(E_{i_1})P(E_{i_2})\dots P(E_{i_k}).$

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▶ Independence implies $P(E_1E_2E_3|E_4E_5E_6) = \frac{P(E_1)P(E_2)P(E_3)P(E_4)P(E_5)P(E_6)}{P(E_4)P(E_5)P(E_6)} = P(E_1E_2E_3)$, and other similar statements.

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- Does pairwise independence imply independence?
- No. Consider these three events: first coin heads, second coin heads, odd number heads. Pairwise independent, not independent.

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- For each a in this countable set, write p(a) := P{X = a}.
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- ▶ Write $F(a) = P\{X \le a\} = \sum_{x \le a} p(x)$. Call *F* the cumulative distribution function.

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- If E_1, E_2, \ldots, E_k are events then $X = \sum_{i=1}^k 1_{E_i}$ is the number of these events that occur.
- Example: in *n*-hat shuffle problem, let *E_i* be the event *i*th person gets own hat.
- Then $\sum_{i=1}^{n} 1_{E_i}$ is total number of people who get own hats.

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Represents weighted average of possible values X can take, each value being weighted by its probability. If the state space S is countable, we can give SUM OVER STATE SPACE definition of expectation:

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Agrees with the SUM OVER POSSIBLE X VALUES definition:

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• How can we compute
$$E[g(X)]$$
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Answer:

$$E[g(X)] = \sum_{x:p(x)>0} g(x)p(x).$$

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- > This is called the **linearity of expectation**.
- Can extend to more variables $E[X_1 + X_2 + \ldots + X_n] = E[X_1] + E[X_2] + \ldots + E[X_n].$

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Defining variance in discrete case

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- ► Taking $g(x) = (x \mu)^2$, and recalling that $E[g(X)] = \sum_{x:p(x)>0} g(x)p(x)$, we find that

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- Variance is one way to measure the amount a random variable "varies" from its mean over successive trials.
- Very important alternate formula: $Var[X] = E[X^2] (E[X])^2$.

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Also, Var[aX] = a²Var[X].

▶ If Y = X + b, where b is constant, then Var[Y] = Var[X].

• Also,
$$\operatorname{Var}[aX] = a^2 \operatorname{Var}[X]$$
.

• Proof: $\operatorname{Var}[aX] = E[a^2X^2] - E[aX]^2 = a^2E[X^2] - a^2E[X]^2 = a^2\operatorname{Var}[X].$

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- If we switch from feet to inches in our "height of randomly chosen person" example, then X, E[X], and SD[X] each get multiplied by 12, but Var[X] gets multiplied by 144.

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- Conclude by additivity of expectation that

$$E[X] = \sum_{j=1}^{n} E[X_j] = \sum_{j=1}^{n} p = np.$$

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• Can show generally that if
$$X_1, \ldots, X_n$$
 independent then $\operatorname{Var}[\sum_{j=1}^n X_j] = \sum_{j=1}^n \operatorname{Var}[X_j]$

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- We think of a Poisson random variable as being (roughly) a Bernoulli (n, p) random variable with n very large and p = λ/n.
- This also suggests $E[X] = np = \lambda$ and $Var[X] = npq \approx \lambda$.

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- The numbers of events occurring in disjoint intervals are independent random variables.
- Probability to see zero events in first t time units is $e^{-\lambda t}$.
- Let T_k be time elapsed, since the previous event, until the *k*th event occurs. Then the T_k are independent random variables, each of which is exponential with parameter λ .

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