# 18.600: Lecture 15 <br> Lectures 1-14 Review 

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## Outline

Counting tricks and basic principles of probability

Discrete random variables

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## Discrete random variables

## Selected counting tricks

- Break "choosing one of the items to be counted" into a sequence of stages so that one always has the same number of choices to make at each stage. Then the total count becomes a product of number of choices available at each stage.


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- How many sequences $a_{1}, \ldots, a_{k}$ of non-negative integers satisfy $a_{1}+a_{2}+\ldots+a_{k}=n$ ?
- Answer: $\binom{n+k-1}{n}$. Represent partition by $k-1$ bars and $n$ stars, e.g., as $* *|* *||* * * *| *$.


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- Countable additivity: $P\left(\cup_{i=1}^{\infty} E_{i}\right)=\sum_{i=1}^{\infty} P\left(E_{i}\right)$ if $E_{i} \cap E_{j}=\emptyset$ for each pair $i$ and $j$.


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- More generally,

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- The notation $\sum_{i_{1}<i_{2}<\ldots<i_{r}}$ means a sum over all of the $\binom{n}{r}$ subsets of size $r$ of the set $\left\{2 \frac{1}{2}, 2, \ldots, n\right\}$.


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- $P\left(\cup_{i=1}^{n} E_{i}\right)=1-\frac{1}{2!}+\frac{1}{3!}-\frac{1}{4!}+\ldots \pm \frac{1}{n!}$
- $1-P\left(\cup_{i=1}^{n} E_{i}\right)=1-1+\frac{1}{2!}-\frac{1}{3!}+\frac{1}{4!}-\ldots \pm \frac{1}{n!} \approx 1 / e \approx .36788$


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- Nice fact: $P\left(E_{1} E_{2} E_{3} \ldots E_{n}\right)=$ $P\left(E_{1}\right) P\left(E_{2} \mid E_{1}\right) P\left(E_{3} \mid E_{1} E_{2}\right) \ldots P\left(E_{n} \mid E_{1} \ldots E_{n-1}\right)$


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- Useful when we think about multi-step experiments.
- For example, let $E_{i}$ be event $i$ th person gets own hat in the $n$-hat shuffle problem.


## Dividing probability into two cases

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\begin{aligned}
P(E) & =P(E F)+P\left(E F^{c}\right) \\
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- In words: want to know the probability of $E$. There are two scenarios $F$ and $F^{c}$. If I know the probabilities of the two scenarios and the probability of $E$ conditioned on each scenario, I can work out the probability of $E$.


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- So $P(A \mid B)$ is $\frac{P(B \mid A)}{P(B)}$ times $P(A)$.
- Ratio $\frac{P(B \mid A)}{P(B)}$ determines "how compelling new evidence is".


## $P(\cdot \mid F)$ is a probability measure

- We can check the probability axioms: $0 \leq P(E \mid F) \leq 1$, $P(S \mid F)=1$, and $P\left(\cup E_{i}\right)=\sum P\left(E_{i} \mid F\right)$, if $i$ ranges over a countable set and the $E_{i}$ are disjoint.


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- The probability measure $P(\cdot \mid F)$ is related to $P(\cdot)$.
- To get former from latter, we set probabilities of elements outside of $F$ to zero and multiply probabilities of events inside of $F$ by $1 / P(F)$.
- $P(\cdot)$ is the prior probability measure and $P(\cdot \mid F)$ is the posterior measure (revised after discovering that $F$ occurs).


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- Equivalent statement: $P(E \mid F)=P(E)$. Also equivalent: $P(F \mid E)=P(F)$.


## Independence of multiple events

- Say $E_{1} \ldots E_{n}$ are independent if for each $\left\{i_{1}, i_{2}, \ldots, i_{k}\right\} \subset\{1,2, \ldots n\}$ we have $P\left(E_{i_{1}} E_{i_{2}} \ldots E_{i_{k}}\right)=P\left(E_{i_{1}}\right) P\left(E_{i_{2}}\right) \ldots P\left(E_{i_{k}}\right)$.


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- In other words, the product rule works.
- Independence implies $P\left(E_{1} E_{2} E_{3} \mid E_{4} E_{5} E_{6}\right)=$ $\frac{P\left(E_{1}\right) P\left(E_{2}\right) P\left(E_{3}\right) P\left(E_{4}\right) P\left(E_{5}\right) P\left(E_{6}\right)}{P\left(E_{4}\right) P\left(E_{5}\right) P\left(E_{6}\right)}=P\left(E_{1} E_{2} E_{3}\right)$, and other similar statements.


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- Does pairwise independence imply independence?
- No. Consider these three events: first coin heads, second coin heads, odd number heads. Pairwise independent, not independent.


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- For each $a$ in this countable set, write $p(a):=P\{X=a\}$. Call $p$ the probability mass function.
- Write $F(a)=P\{X \leq a\}=\sum_{x \leq a} p(x)$. Call $F$ the cumulative distribution function.


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- Example: in $n$-hat shuffle problem, let $E_{i}$ be the event $i$ th person gets own hat.
- Then $\sum_{i=1}^{n} 1_{E_{i}}$ is total number of people who get own hats.


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- Represents weighted average of possible values $X$ can take, each value being weighted by its probability.


## Expectation when state space is countable

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- Agrees with the SUM OVER POSSIBLE $X$ VALUES definition:

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- This is called the linearity of expectation.
- Can extend to more variables

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E\left[X_{1}+X_{2}+\ldots+X_{n}\right]=E\left[X_{1}\right]+E\left[X_{2}\right]+\ldots+E\left[X_{n}\right]
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- Very important alternate formula: $\operatorname{Var}[X]=E\left[X^{2}\right]-(E[X])^{2}$.


## Identity

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- Also, $\operatorname{Var}[a X]=a^{2} \operatorname{Var}[X]$.
- Proof: $\operatorname{Var}[a X]=E\left[a^{2} X^{2}\right]-E[a X]^{2}=a^{2} E\left[X^{2}\right]-a^{2} E[X]^{2}=$ $a^{2} \operatorname{Var}[X]$.


## Standard deviation

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- Uses the same units as $X$ itself.
- If we switch from feet to inches in our "height of randomly chosen person" example, then $X, E[X]$, and $\mathrm{SD}[X]$ each get multiplied by 12 , but $\operatorname{Var}[X]$ gets multiplied by 144 .


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- Number of heads is binomial random variable with parameters $(n, p)$.


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- Note that $E\left[X_{j}\right]=p \cdot 1+(1-p) \cdot 0=p$ for each $j$.
- Conclude by additivity of expectation that

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## Compute variance with decomposition trick

- $X=\sum_{j=1}^{n} X_{j}$, so
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- Can show generally that if $X_{1}, \ldots, X_{n}$ independent then $\operatorname{Var}\left[\sum_{j=1}^{n} X_{j}\right]=\sum_{j=1}^{n} \operatorname{Var}\left[X_{j}\right]$


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- We think of a Poisson random variable as being (roughly) a Bernoulli ( $n, p$ ) random variable with $n$ very large and $p=\lambda / n$.
- This also suggests $E[X]=n p=\lambda$ and $\operatorname{Var}[X]=n p q \approx \lambda$.


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- The numbers of events occurring in disjoint intervals are independent random variables.
- Probability to see zero events in first $t$ time units is $e^{-\lambda t}$.
- Let $T_{k}$ be time elapsed, since the previous event, until the $k$ th event occurs. Then the $T_{k}$ are independent random variables, each of which is exponential with parameter $\lambda$.


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### 18.600 Probability and Random Variables

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