## Lecture 16

## Statistical Analysis in Biomaterials Research (Part II)

## C. F Distribution

$>$ Allows comparison of variability of behavior between populations using test of hypothesis: $\sigma_{x}=\sigma_{x}$,

Define a statistic:

Named for British statistician
Sir Ronald A. Fisher.
$\chi_{v_{1}}{ }^{2}=v_{1} S^{2} / \sigma^{2}$ where $v_{1}$ is degrees of freedom.
then $F=\frac{\chi_{v_{1}}{ }^{2} / v_{1}}{\chi_{v_{2}}{ }^{2} / v_{2}}$

For $\sigma_{x}=\sigma_{x^{\prime}} \Rightarrow F=\frac{S_{x}^{2}}{S_{x^{\prime}}^{2}}$


Procedure to test variability hypothesis:

1. Calculate $S_{x}{ }^{2}$ and $S_{X},{ }^{2}$ (with $v_{1}=N-1$ and $v_{2}=N^{\prime}-1$, respectively)
2. Compute F
3. Look in $F$-distribution tables for critical $F$ for $\nu_{1}, \nu_{2}$, and desired confidence level $P$
4. For $F_{\frac{1-P}{2}}<F<F_{\frac{1+P}{2}} \Rightarrow \sigma_{\mathrm{x}}=\sigma_{\mathrm{x}}$,

Case Example: Measurements of C5a production for blood exposure to an extracorpeal filtration device and tubing gave same means, but different variabilities. Are the standard deviations different within 95\% confidence?

Control (tubing only): $\mathrm{S}_{\mathrm{x}}{ }^{2}=26(\mu \mathrm{~g} / \mathrm{ml})^{2}, \mathrm{v}_{2}=9$
Filtration device: $\mathrm{S}_{\mathrm{x}}{ }^{2}=32(\mu \mathrm{~g} / \mathrm{ml})^{2}, v_{1}=7$

1. Calculate $S_{x}{ }^{2}$ and $S_{x}{ }^{2}{ }^{2}$ (provided)
2. Compute F

$$
F=\frac{S_{x}^{2}}{S_{x^{\prime}}^{2}}=32 / 26=1.231
$$

3. Determine critical $F$ values from $F$-distribution chart
$v_{1}=7$ and $v_{2}=9(\mathrm{~m}, \mathrm{n}$ for use with tables)

$$
\begin{aligned}
& \frac{1-P}{2}=0.025 \Rightarrow F_{0.025}=0.207 \\
& \frac{1+P}{2}=0.975 \Rightarrow F_{0.975}=4.20
\end{aligned}
$$

For $0.207 \leq F \leq 4.20 \Rightarrow \sigma_{x}=\sigma_{x}$,
$F=1.231$ falls within this interval.
Conclude $\sigma$ values for two systems are the same!

## D. Other distributions of interest

$>$ Radioactive decay $\Rightarrow$ Poisson distribution
$>$ Only 2 possible outcomes $\Rightarrow$ Binomial distribution

## 3. Standard deviations of computed values

$>$ If quantity $z$ of interest is a function of measured parameters

$$
\mathrm{z}=\mathrm{f}(\mathrm{x}, \mathrm{y}, \ldots)
$$

What is $\mathrm{S}_{\mathrm{z}}$ ?

$$
\text { We assume: }\langle z\rangle=f(\langle x\rangle,\langle y\rangle, \ldots)
$$

Deviations ( $\delta z$ ) of $z$ from its universal value can be written as:

$$
\partial z=\frac{\partial z}{\partial x} \partial x+\frac{\partial z}{\partial y} \partial y+\ldots
$$

The standard deviation for z is calculated:

$$
S_{z}=\sqrt{\left(\frac{\partial\langle z\rangle}{\partial\langle x\rangle}\right)^{2} S_{x}^{2}+\left(\frac{\partial\langle z\rangle}{\partial\langle y\rangle}\right)^{2} S_{y}^{2}+\ldots}
$$

Case Example: We measure motility $(\mu)$ and persistence $(P)$ of a cell and want to know the standard deviation of the speed:

$$
\langle\mu\rangle=\frac{\langle S\rangle^{2}\langle P\rangle}{2} \text { rearranges to }\langle S\rangle=\sqrt{\frac{2\langle\mu\rangle}{\langle P\rangle}}
$$

$$
\frac{\partial\langle S\rangle}{\partial\langle P\rangle}=-0.5 \sqrt{2\langle\mu\rangle}\langle P\rangle^{-3 / 2}=\frac{-\langle S\rangle}{2\langle P\rangle} \quad \frac{\partial\langle S\rangle}{\partial\langle\mu\rangle}=0.5 \frac{\sqrt{2}}{\langle P\rangle\langle\mu\rangle}=\frac{\langle S\rangle}{2\langle\mu\rangle}
$$

$$
S_{S}=\sqrt{\left(\frac{\partial\langle S\rangle}{\partial\langle P\rangle}\right)^{2} S_{P}^{2}+\left(\frac{\partial\langle S\rangle}{\partial\langle\mu\rangle}\right)^{2} S_{\mu}^{2}}=\sqrt{\frac{\langle S\rangle^{2}}{4\langle P\rangle^{2}} S_{P}^{2}+\frac{\langle S\rangle^{2}}{4\langle\mu\rangle^{2}} S_{\mu}^{2}}
$$

## 4. Least Squares Analysis of Data (Linear Regression)

$>$ Computing the straight line that best fits data.
Suppose we have some measured data for binding of a ligand to its receptor:


This equilibrium is described by: $\mathrm{K}=[\mathrm{C}] /[\mathrm{L}][\mathrm{R}]$
$v=$ fraction of occupied receptors $=[\mathrm{C}] /([\mathrm{C}]+[\mathrm{R}])=\mathrm{K}[\mathrm{L}] /(1+\mathrm{K}[\mathrm{L}])$
$1 / v=1+1 / \mathrm{K}[\mathrm{L}]$

Question: How can we numerically obtain the linear equation that best represents the data?

Answer: Minimize the squared deviation of the line from each point.

| NOTE: This is a |
| :--- |
| generic tool in |
| data regression, |
| independent of the |
| fitting function. |



The deviation of any given measured point $\left(\mathrm{x}_{\mathrm{i}}, \mathrm{y}_{\mathrm{i}}\right)$ from the line is:

$$
\mathrm{y}_{\mathrm{i}}-\mathrm{y}_{\text {line }}=\mathrm{y}_{\mathrm{i}}-\left(\mathrm{mx}_{\mathrm{i}}+\mathrm{b}\right)
$$

Where $m$ and $b$ are the slope and intercept of the line.
Our minimization criterion can thus be written:

$$
M=\sum_{i=1}^{N}\left[y_{i}-\left(m x_{i}+b\right)\right]^{2}=\text { minimum }
$$

Mathematically we require: $\frac{\partial M}{\partial m}=0, \frac{\partial M}{\partial b}=0$

Solving these two equations for the two unknowns (best fit $m$ and $b$ for the line), we get:

$$
m=\frac{N \sum_{i=1}^{N}\left(x_{i} y_{i}\right)-\sum_{i=1}^{N} x_{i} \sum_{i=1}^{N} y_{i}}{N \sum_{i=1}^{N} x_{i}^{2}-\left(\sum_{i=1}^{N} x_{i}\right)^{2}} \quad b=\frac{\sum_{i=1}^{N} y_{i}-m \sum_{i=1}^{N} x_{i}}{N}
$$

## $>$ Quantifying Error of the Straight-Line Fit

If the error on each $y_{i}$ is unknown (e.g., a single measurement was made):
The standard deviation for the regression line is given by:

This assumes:

$$
\sigma=\sqrt{\frac{M}{N-2}} \quad \begin{aligned}
& \begin{array}{l}
N-2 \text { in denominator since 2 degrees of } \\
\text { freedom are taken in calculating } \mathrm{m} \text { and } \\
\mathrm{b} \text { (two points make a line, so } \sigma \text { for } N=2 \\
\text { is meaningless.) }
\end{array} \\
& \hline
\end{aligned}
$$

- a normal distribution of data points about the line
- spread of points is of similar magnitude for full data range
"Goodness" of fit can be further characterized by the correlation coefficient, $r$ (or coefficient of determination, $r^{2}$ ), calculated as:

$$
r^{2}=\frac{\sum_{i=1}^{N}\left(y_{i}-\langle y\rangle\right)^{2}-M}{\sum_{i=1}^{N}\left(y_{i}-\langle y\rangle\right)^{2}} \quad \begin{aligned}
& \text { For a "perfect fit" } \\
& M=0 \Rightarrow r^{2}=1
\end{aligned}
$$

For $r^{2}=1,<y>$
represents data as
well as a line
$>$ Many calculators, spreadsheets \& other math tools are programmed to perform linear least-squares fitting, as well as fits to more complex equations following a similar premise.
$>$ Many nonlinear equations can be linearized by taking the log of both sides
e.g.,

$$
y=b x^{m} \text { becomes } \ln y=m \ln x+\ln b
$$

$$
\text { or } \quad y^{\prime}=m x^{\prime}+b^{\prime}
$$

> Multiple regression
In some cases, we wish to fit data dependent on more than one independent variable. The procedure will be exactly analogous to that used above, and solutions can be obtained through matrix algebra.

Here we will consider the simple case of a linear dependence on 2 independent variables.

Our minimization criterion can thus be written:

$$
M=\sum_{i=1}^{N}\left[y_{i}-\left(a+b x_{i}+c z_{i}\right)\right]^{2}=\text { minimum }
$$

Mathematically we require: $\frac{\partial M}{\partial a}=0, \frac{\partial M}{\partial b}=0, \frac{\partial M}{\partial c}=0$
which yields the 3 equations:

$$
\begin{aligned}
N a+\left(\sum_{i=1}^{N} x_{i}\right) b+\left(\sum_{i=1}^{N} z_{i}\right) c & =\sum_{i=1}^{N} y_{i} \\
\left(\sum_{i=1}^{N} x_{i}\right) a+\left(\sum_{i=1}^{N} x_{i}^{2}\right) b+\left(\sum_{i=1}^{N}\left(x_{i} z_{i}\right)\right) c & =\sum_{i=1}^{N}\left(x_{i} y_{i}\right) \\
\left(\sum_{i=1}^{N} z_{i}\right) a+\left(\sum_{i=1}^{N} x_{i} z_{i}\right) b+\left(\sum_{i=1}^{N}\left(z_{i}^{2}\right)\right) c & =\sum_{i=1}^{N}\left(z_{i} y_{i}\right)
\end{aligned}
$$

These equations can be solved to obtain $a, b$ and $c$.

## References

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