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# **Martingales II**

#### Content.

- 1. Review
- 2. Some applications of Optional Stopping Theorem
- 3. Martingale Convergence Theorem

# 1 Review

**Definition 1** (Martingale).  $\{M_t\}$  is a martingale with respect to  $\mathcal{F}_0 \subset \mathcal{F}_1 \subset \ldots \subset \mathcal{F}$  if it satisfies: 1.  $M_t \in \mathcal{F}_t, t \ge 0$ 2.  $\mathbb{E}|M_t| < \infty, t \ge 0$ 3.  $\mathbb{E}[M_t|\mathcal{F}_{t-1}] = M_{t-1}, t \ge 1$ 

In other words,  $\{M_t\}$  is a martingale w.r.t.  $\{X_t\}$  if:

- 1.  $M_t = f(X_0, ..., X_t), t \ge 0$ 2.  $\mathbb{E}|M_t| < \infty, t \ge 0$
- 3.  $\mathbb{E}[M_t|X_0,\ldots,X_{t-1}] = M_{t-1}, t \ge 1.$

For convenience, we denote it by

$$\mathbb{E}_s = \mathbb{E}[\cdot | \mathcal{F}_s] = \mathbb{E}[\cdot | X_0, \dots, X_s],$$

and the third condition can be written as:  $\mathbb{E}_{t-1}M_t = M_{t-1}$  for any  $t \ge 1$ .

When we say  $\{M_t\}$  is a martingale without specifying the filtration, we mean that it is a martingale with respect to its natural filtration, i.e.  $\mathcal{F}_t = \sigma(M_0, \dots, M_t)$ . We consider it as a special case of the definition.

Now if  $\{M_t\}$  is a martingale with respect to the filtration  $\mathcal{F}_t = \sigma(X_0, \ldots, X_t)$ , it is also a martingale with respect to its filtration  $\sigma(M_0,\ldots,M_t)$ . In fact, by the tower property, we have:

 $\mathbb{E}[M_t|M_0,\ldots,M_{t-1}] = \mathbb{E}[\mathbb{E}[M_t|\mathcal{F}_t]|M_0,\ldots,M_{t-1}] = \mathbb{E}[M_{t-1}|M_0,\ldots,M_{t-1}] = M_{t-1}.$ 

**Properties of a martingale**  $\{M_t\}$ 

- E<sub>s</sub>M<sub>t</sub> = M<sub>t∧s</sub>
  EM<sub>t</sub> = EM<sub>0</sub>, ∀t ≥ 0
  If M<sub>t</sub> is a martingale, and τ is a stopping time, then Y<sub>t</sub> = M<sub>t∧τ</sub> is a martingale.
- Side note:  $\mathbb{E}_n$  also works like a martingale:  $\mathbb{E}_n \mathbb{E}_m = \mathbb{E}_{n \wedge m}$ ,  $\mathbb{E} \mathbb{E}_n =$  $\mathbb{E}_n\mathbb{E} = \mathbb{E}$ . In fact, you can also define  $\mathbb{E}_{\tau}$  and even have  $\mathbb{E}_{\tau}M_t = M_{\tau \wedge t}$ . But we won't do it in this class.

Let  $A_t$  be the gambler's ruin Markov chain starting from k. Now let's consider the simple random walk  $S_t$  starting from  $S_t = k$ , and  $S_t = S_{t-1} + X_t$  with  $\mathbb{P}(X_t = \pm 1) = \frac{1}{2}$ . Let,

$$\tau = \inf\{t : S_t = 0 \text{ or } S_t = n\}.$$

Notice that  $A_t = S_{t \wedge \tau}$ . Since  $S_t$  is a martingale, it follows that its stopped martingale  $A_t$  is a martingale as well. This implies a good property of gambler's ruin Markov chain, which is

$$\mathbb{E}A_t = A_0 = k.$$

We will use the definitions and notations of  $S_t$  and  $A_t$  for several times in this lecture.

**Theorem 1** (Optional Stopping Theorem). If  $\{M_t\}$  is a martingale and  $\tau$  is a stopping time such that  $\{M_t\}$  is uniformly integrable and  $\mathbb{P}(\tau < \infty) = 1$ , then

$$\mathbb{E}M_{\tau} = \mathbb{E}M_0.$$

**Proposition 1** (Uniformly integrable martingales). *The following propositions about uniformly integrable martingales (u.i.M.) hold:* 

- 1.  $M_t = \mathbb{E}[Z|\mathcal{F}_t]$  for any Z such that  $\mathbb{E}|Z| < \infty$  is always u.i.M.
- 2. If there exists G(t) such that  $G(t)/t \to \infty$  as  $t \to \infty$ . If  $\sup_t \mathbb{E}[G(M_t)] < \infty$ , then  $M_t$  is u.i.M.
- 3. If  $|M_t M_{t-1}| \le c < \infty$  and  $\mathbb{E}\tau < \infty$ , then  $Y_t = M_{t \wedge \tau}$  is u.i.M.

 $\{S_n\}$  is not uniformly integrable. Indeed, the magnitude of  $|S_n|$  is approximately  $O(\sqrt{n})$ . For any b, one can always find some  $N \approx b^2$  such that  $\mathbb{E}[|S_N|1\{|S_N| \ge b\}] \ge c$ , so  $\sup_n \mathbb{E}[|S_n|1\{|S_n| \ge b\}] \not\to 0$  as  $b \to \infty$ . Hence,  $S_n$  is not uniformly integrable.

# 2 Some applications of O.S.T.

#### 2.1 Gambler's Ruin

For the **gambler's ruin** problem, we start with  $A_0 = k$ , and we want to find  $\mathbb{P}[win] = \mathbb{P}[A_{\infty} = n]$ .

Note that  $A_t = S_{t \wedge \tau}$  is a u.i.M (since  $A_t$  is bounded), therefore

$$n\mathbb{P}[win] = \mathbb{E}A_{\tau} = \mathbb{E}A_0 = k_1$$

as

$$A_{\tau} = \begin{cases} 0, & \text{"ruined"} \\ n, & \text{"won"} \end{cases}$$

Therefore,  $\mathbb{P}[win] = \frac{k}{n}$ .

Now let  $M_t = S_t^2 - t = (S_{t-1} + X_t)^2 - t = S_{t-1}^2 + 2X_tS_{t-1} + X_t^2 - t = S_{t-1}^2 - (t-1) + 2X_tS_{t-1} = M_{t-1} + 2X_tS_{t-1}$ . Therefore,

$$\mathbb{E}_{t-1}M_t = M_{t-1}.$$

 $M_{t\wedge\tau}$  is uniformly integrable since the increment  $|M_t - M_{t-1}| = 2|X_tS_{t-1}|$  is bounded.

Therefore, by OST, we have

$$\frac{k}{n} \cdot n^2 - \mathbb{E}\tau = \mathbb{E}[M_\tau] = M_0 = k^2,$$

and thus,  $\mathbb{E}\tau = k(n-k)$ .

#### **2.2** Null recurrence of $S_t$

We start with  $S_0 = k$ . Let  $\tau_1 = \inf\{t : S_t = 0\}$ , and  $B_t = S_{t \wedge \tau_1}$ . One can think of  $B_t$  as a Markov chain with 0 the absorbing state.

We know from recurrence of  $S_t$  that  $\tau_1 < \infty$  a.s.. We also know that  $\mathbb{E}S_t = \mathbb{E}B_t = \mathbb{E}S_0 = k$ .

If  $B_t$  were a u.i.M, then OST applies, we will have  $\mathbb{E}B_{\tau_1} = k$ . However, by definition,  $B_{\tau_1} = 0$  a.s., so  $\mathbb{E}B_{\tau_1} = 0 \neq k$ . By Proposition 2(3), the only thing that prevents  $B_t$  from being a u.i.M is  $\mathbb{E}\tau = \infty$ . Therefore,  $S_t$  is null recurrent.

#### 2.3 Gambler's Ruin in the asymmetric case

For the asymmetric case, i.e.  $S_t = S_{t-1} + X_t$  with  $\mathbb{P}(X_t = 1) = p$  and  $\mathbb{P}(X_t = -1) = 1 - p$ , one can use the following two martingales to compute  $\mathbb{P}[win]$  and  $\mathbb{E}\tau$ :

1.  $M_t = S_t - (2p - 1)t$ 

2. 
$$N_t = e^{\lambda S_t - t\psi_X(\lambda)}$$
, where  $\psi_{X_1}(\lambda) = \ln M_{X_1}(\lambda)$ 

From OST, we have

$$\mathbb{E}S_{\tau} - (2p-1)\mathbb{E}\tau = \mathbb{E}M_{\tau} = M_0 = k$$

and

$$e^{\lambda n} \mathbb{P}[win] + \mathbb{P}[ruined] = \mathbb{E}N_{\tau} = N_0 = e^{\lambda k},$$

with some  $\lambda (= \ln \frac{p}{1-p})$  such that  $\psi_{X_1}(\lambda) = 0$ .

#### 3 Martingale Convergence Theorem

Think of  $M_t$  as the price of stock. At time t - 1, you decide to move your possession of stock to  $F_t$  shares, where  $F_t \in \mathcal{F}_{t-1}$  is determined by all the

observed information at time t - 1. Then the value of your portfolio at time t is

$$V_t = F_0 M_0 + F_1 (M_1 - M_0) + \ldots + F_t (M_t - M_{t-1}) \stackrel{\triangle}{=} \int_0^t F \, dM.$$

**Proposition 2.** If  $M_t$  is a martingale, then  $V_t$  is a martingale. In particular,  $\mathbb{E}V_t = \mathbb{E}V_0$ .

The important consequence is that if you start with  $F_0$  shares priced at  $M_0$  then no trading strategy (and no finite cash-out time) can yield an expectation different from what you had  $\mathbb{E}[F_0M_0]$  in the beginning. Assuming the market price is a martingale with respect to the same filtration  $\mathcal{F}_t$  that determines the available information you have to execute the trading decisions.

**Definition 2.** Starting  $S_0 = 0$ , define  $T_k = \inf\{t \ge S_{k-1} : M_t \le a\}$ ,  $S_k = \inf\{t \ge T_k : M_t \ge b\}$ . Define  $U_n(a, b) =$ # of upcrossings of (a, b) in  $0 \le t \le n$ , i.e.

$$U_n(a,b) = \sup\{k : S_k \le n\}.$$

Lemma 1 (Upcrossing Lemma).

$$\mathbb{E}[U_n(a,b)] \le \frac{\mathbb{E}(M_n-a)_{-}}{b-a}.$$

*Proof.* Starting with  $F_0 = 0$  and do trading: buy 1 share when  $M_t \le a$  and sell it when  $M_t \ge b$ . Since  $V_0 = 0$ , we have

$$V_n \ge (b-a)U_n + (M_n - a) \land 0 = (b-a)U_n - (M_n - a)_{-1}$$

Since  $V_n$  is a martingale, it follows from Optional Stopping Theorem that

$$\mathbb{E}U_n \le \frac{\mathbb{E}(M_n - a)_{-}}{b - a}.$$

**Theorem 2.** If  $M_n$  is a martingale such that  $\mu \quad \sup_n \mathbb{E}|M_n| < \infty$ , then there exists an integrable random variable  $M_\infty$  such that

$$M_n \stackrel{\text{a.s.}}{\to} M_\infty$$
, and  $\mathbb{E}[|M_\infty|] \le \mu < \infty$ .

If  $M_t$  is u.i.M, then  $M_t \stackrel{L_1}{\rightarrow} M_\infty$  and

$$M_t = \mathbb{E}[M_\infty | \mathcal{F}_t].$$

Remark: Note that if  $M_t$  is u.i.M. then  $\mu < \infty$  automatically. Thus, the second part of the theorem shows that every u.i.M. is in fact a Doob martingale.

*Proof.* **Proof of part 1**: Fix b > a,

$$U(a,b) = \lim_{n \to \infty} U_n(a,b).$$

By the upcrossing lemma, we have

$$\mathbb{E}U_n(a,b) \le \frac{\mathbb{E}(M_n - a)_{-}}{b - a} \le \sup_n \frac{\mathbb{E}|M_n| + |a|}{b - a} < \infty.$$

Therefore, by Monotone Convergence Theorem, we have

$$\mathbb{E}U(a,b) = \lim_{n \to \infty} \mathbb{E}U_n(a,b) \le \sup_n \frac{\mathbb{E}|M_n| + |a|}{b-a} < \infty.$$

This imples that,

$$\mathbb{P}(U(a,b) = \infty \text{ for any } b > a, a, b \in \mathbb{Q}) = 0.$$

So with probability 1 the trajectory  $M_n$  intersects any arbitrary small interval only finitely many times. Thus there must exist a (possibly extended real-valued) random variable  $M_\infty$  such that  $M_n \xrightarrow{a.s.} M_\infty$ .

To show that  $M_{\infty}$  is in fact integrable (and hence real-valued) we use Fatou's lemma:

$$\mathbb{E}[|M_{\infty}|] = \mathbb{E}[\liminf_{n \to \infty} |M_n|] \le \liminf_{n \to \infty} \mathbb{E}[|M_n|] \le \mu < \infty$$

**Proof of part 2:** To show  $M_t = \mathbb{E}[M_{\infty}|\mathcal{F}_t]$ , it suffices to show that for any  $B \in \mathcal{F}_t$ , we have

$$\mathbb{E}M_{\infty}\mathbf{1}_B = \mathbb{E}M_t\mathbf{1}_B.$$

For any  $m \ge t$ , we have

$$\mathbb{E}M_m 1_B = \mathbb{E}[\mathbb{E}_t[M_m 1_B]] = \mathbb{E}[1_B M_t].$$

Since  $M_m 1_B \xrightarrow{a.s.} M_\infty 1_B$  and  $\{M_m 1_B\}$  is uniformly integrable, it follows that  $M_m 1_B \xrightarrow{L_1} M_\infty 1_B$ . Therefore,

$$\mathbb{E}[M_t 1_B] = \lim_{m \to \infty} \mathbb{E}[M_m 1_B] = \mathbb{E}[M_\infty 1_B].$$

**Corollary 1.** If  $M_n \ge 0$ ,  $M_n$  is a martingale, then it converges almost surely to integrable  $M_{\infty}$ .

*Proof.* Since for any n,

$$\mathbb{E}|M_n| = \mathbb{E}M_n = \mathbb{E}M_0,$$

it follows that

$$\sup_{n} \mathbb{E}|M_n| < \infty.$$

In particular,  $M_n = X_1 \dots X_n$  such that  $X_n \ge 0, \mathbb{E}X_n = 1$ . Then,  $M_n$  converges almost surely.

## 4 Further topics

Martingale and stopping time theory is rich subject. The key omissions are:

- A lot of results about martingales are also available for submartingales (i.e. when  $\mathbb{E}_{t-1}[M_t] \ge M_{t-1}$ ) and supermartingales (i.e. when  $\mathbb{E}_{t-1}[M_t] \le M_{t-1}$ ).
- Maximal inequalities for martingales/submartingales/supermartingales). These establish results similar to Kolmogorov's maximal inequalities (for sums of independent r.v.s) but for general martingales. To get a flavor of such results, if  $M_0 = 0$  then

$$\mathbb{P}[\max_{0 \le t \le n} M_t > b] = \mathbb{P}[U_n(0, b) \ge 1] \le \frac{1}{b} \mathbb{E}[|M_n|],$$

where in the last step we applied the upcrossing Lemma and Markov's inequality. So in particular, in the setting of convergence theorem we see that life-time maximum of  $M_t$  is of the order of  $\mu$ . Other maximal inequalities bound *p*-th norm of the maximum in terms of the *p*-th norm of  $M_n$  etc.

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