7.90J 6.874J Computational functional genomics (Spring 2005: Lecture 13)

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Topics

- Modeling Biological Systems
 - A simple biological system
 - Model assumptions
- Discrete Bayesian networks
 - Discretizing data
 - Bayesian scoring functions
 - Edge scores

A simple biological system

- Data are observed from a biological subnetwork with four genes
- The genes might influence one anothers expression
- The structure of the network is hidden from us (we do not see the edges)
- We observe mRNA and active protein levels for each of the four genes
- We have hundreds of observations

Model assumptions

- There are no hidden variables (only A, B, C, and D can influence one another)
- There are no cycles in the unknown network
- "Sufficient conditions" are observed to perturb the expession of A, B, C, and D
- All observed data are continuous
- Data is complete (no missing variable observations)
- The underlying biological system can be modeled using discrete states
- Uniform population behavior (Why is this important?)
- We begin with 8 nodes...

Bayesian networks

- Nodes represent variables
- Each node has 0 or more parents
- The structure S of edges describes how the joint probability distribution of the observed variables can be factored
- $\bullet~S$ encodes the conditional independence of the observed variables
- To fully specify a network we need to specify how children depend on their parents
- This dependency is encoded in the parameters θ
- Given *n* variables, roughly how many structures are there?
- Less than $(2^n)^n = 2^{n^2}$ (Great!)

Bayesian network tasks

- Learn the structure of a Bayesian network (S) given observed data (Structure Learning)
- Learn a Bayesian network (S and θ) given observed data (Learning)
- Infer X_j when it is not observed given S and θ (Inference)

Discrete Bayesian Networks - Interval discretization

- Sort the observed values from smallest to largest
- Divide range of observed values into L intervals
- Policy vector

$$\Lambda = (-\infty, x_0 + \frac{(x_{N-1} - x_0)}{L}, x_0 + \frac{2(x_{N-1} - x_0)}{L}, \dots,$$
(1)
$$x_0 + \frac{(L-1)(x_{N-1} - x_0)}{L}, \infty)$$
(2)

Quantile discretization

- Place an equal number of observations into L levels
- Policy vector

$$\Lambda = (-\infty, \frac{x \lfloor \frac{N}{L} \rfloor + x \lfloor \frac{N}{L} \rfloor + 1}{2}, \frac{x \lfloor \frac{2N}{L} \rfloor + x \lfloor \frac{2N}{L} \rfloor + 1}{2}, \dots, \qquad (3)$$
$$\frac{x \lfloor \frac{(L-1)N}{L} \rfloor + x \lfloor \frac{(L-1)N}{L} \rfloor + 1}{2}, \infty) \qquad (4)$$

How do we decide on the number of levels?

- We can begin with one level for each unique observed value
- If we start with L levels of discretization, we can reduce this to L-1 by coalescing levels
- Coalese two levels by adding the probabilities of the merged levels
- For example, we could start with 10 levels for 10 observations, and then reduce this to L = 1



How should we merge levels?

• We could consider variables independently





How should we merge levels?

• Or consider preserving information between genes



Total mutual information between variables

- Let vector X_i^L be the discretization of variable X_i from all observations into L levels
- Define the total mutual information between all X_i^L at discretization level L as:

$$TMI(L) = \sum_{i,j} H(X_i^L) + H(X_j^L) - H(X_i^L, X_j^L)$$
(5)

- Mutual information is 0 when variables are independent
- $H(X_i^L)$ is a measure of the randomness of X_i^L

$$H(X_i^L) = -\sum_{X_i^L} p(X_i^L) log(p(X_i^L))$$
(6)

• $H(X_i^L, X_j^L)$ is the mutual entropy of X_i^L and X_j^L

$$H(X_{i}^{L}, X_{j}^{L}) = -\sum_{X_{i}^{L}, X_{j}^{L}} p(X_{i}^{L}, X_{j}^{L}) log(p(X_{i}^{L}, X_{j}^{L}))$$
(7)

Total mutual information as a function of L

- When we go from L to L-1, pick the levels to merge to minimize TMI(L) TMI(L-1)
- As we decrease L, TMI(L) decreases:



- Pick an L that captures most of the information
- Why do we want to reduce L?

The Bayesian scoring metric

• The Bayesian score of model S given observed data D can be decomposed into a likelihood and a prior

$$BayesianScore(S) = \log p(S|D)$$
(8)

$$= \log p(S) + \log p(D|S) + c$$
 (9)

• The likelihood function is computed as follows

$$p(D|S) = \int_{\theta} p(D, \theta|S) d\theta$$
 (10)

$$= \int_{\theta} p(D|\theta, S) p(\theta|S) d\theta$$
(11)

Parameters for discrete Bayesian networks

- \bullet Index the n variables in the Bayesian network using the variable i
- Index the q_i parent configurations of variable i using the variable j
- Index the r_i states of variable i using the variable k
- θ_{ijk} is the probability of observing variable i in state k given parent configuration j

$$(\theta_{ij1}, \dots, \theta_{ijr_i}) \sim \text{Dirichlet}(\alpha_{ij1}, \dots, \alpha_{ijr_i}) \quad \forall i, j \quad (12)$$

$$\sim c \cdot \theta_{ij1}^{\alpha_{ij1}-1} \theta_{ij2}^{\alpha_{ij2}-1} \dots \theta_{ijr_i}^{\alpha_{ijr_i}-1}$$
(13)

• α are the hyperparameters

Scoring discrete Bayesian networks

- Assign each observation to a single level
- Let N_{ijk} be the number of occurrences in the data set D of variable i in state k given parent configuration j and

$$N_{ij} = \sum_{k=1}^{r_i} N_{ijk}$$
(14)
$$\alpha_{ij} = \sum_{k=1}^{r_i} \alpha_{ijk}$$
(15)

• The Bayesian score of S (see Heckerman on the Web site) is:

$$\log p(S) + \log \left\{ \prod_{i=1}^{n} \prod_{j=1}^{q_i} \left(\frac{\Gamma(\alpha_{ij})}{\Gamma(\alpha_{ij} + N_{ij})} \cdot \prod_{k=1}^{r_i} \frac{\Gamma(\alpha_{ijk} + N_{ijk})}{\Gamma(\alpha_{ijk})} \right) \right\} (16)$$
$$\log p(S) + \sum_{i=1}^{n} \sum_{j=1}^{q_i} \left\{ \log \frac{\Gamma(\alpha_{ij})}{\Gamma(\alpha_{ij} + N_{ij})} + \sum_{k=1}^{r_i} \log \frac{\Gamma(\alpha_{ijk} + N_{ijk})}{\Gamma(\alpha_{ijk})} \right\} (17)$$

Example: Yeast pheromone response pathway

Image removed for copyright reasons.

Total mutual information as a function of L

• We start with 320 experiments with L = 160 and run the level merging algorithm that minimizes the loss in total mutual information



• L = 4 for scoring

Model search: distribution of model scores

Histogram of model scores using a random walk and simulated annealing. Note that simulated annealing does not get stuck as easily



Model averaging

- Integrating over all possible parameters protects us from overfitting parameters
- We can provide some protection against overfitting model structure by averaging over the model posterior distribution

$$p(E_{XY}|D) = \sum_{S} p(E_{XY}, S|D)$$
(18)

$$= \sum_{S} p(E_{XY}|D,S) \cdot p(S|D)$$
(19)

$$= \sum_{S} \mathbf{1}_{XY}(S) \cdot e^{BayesianScore(S)}$$
(20)

Model search: edge consensus of top 50 models

Edge colors: black 10^9 , purple $10^6 - 10^9$, dark blue $10^3 - 10^6$, light blue $1 - 10^3$



Model search: edge consensus of top 50 constrained models



How can we model these?

turns on lowers represses deinhibits methylates dephosphorylates reduces translates catalyzes binds silences promotes is necessary for is a factor in

turns off activates derepresses expresses demethylates protects oxidizes regulates metabolizes initiates stimulates requires is a component of

raises deactivates inhibits suppresses phosphorylates deprotects transcribes controls ligates enhances induces elevates is a substitute for

Idea - use the parameter prior!

• Recall the likelihood function is:

$$p(D|S) = \int_{\theta} p(D, \theta|S) d\theta$$
 (21)

$$= \int_{\theta} p(D|\theta, S) p(\theta|S) d\theta$$
 (22)

- We can use $p(\theta|S)$ to model relationships
- For example, to represent a positive edge from X to Y, for all values y of Y, and for all values $x_i < x_j$ of X we constrain θ so that:

$$p(Y > y | X = x_i) \le p(Y > y | X = x_j)$$
 (23)