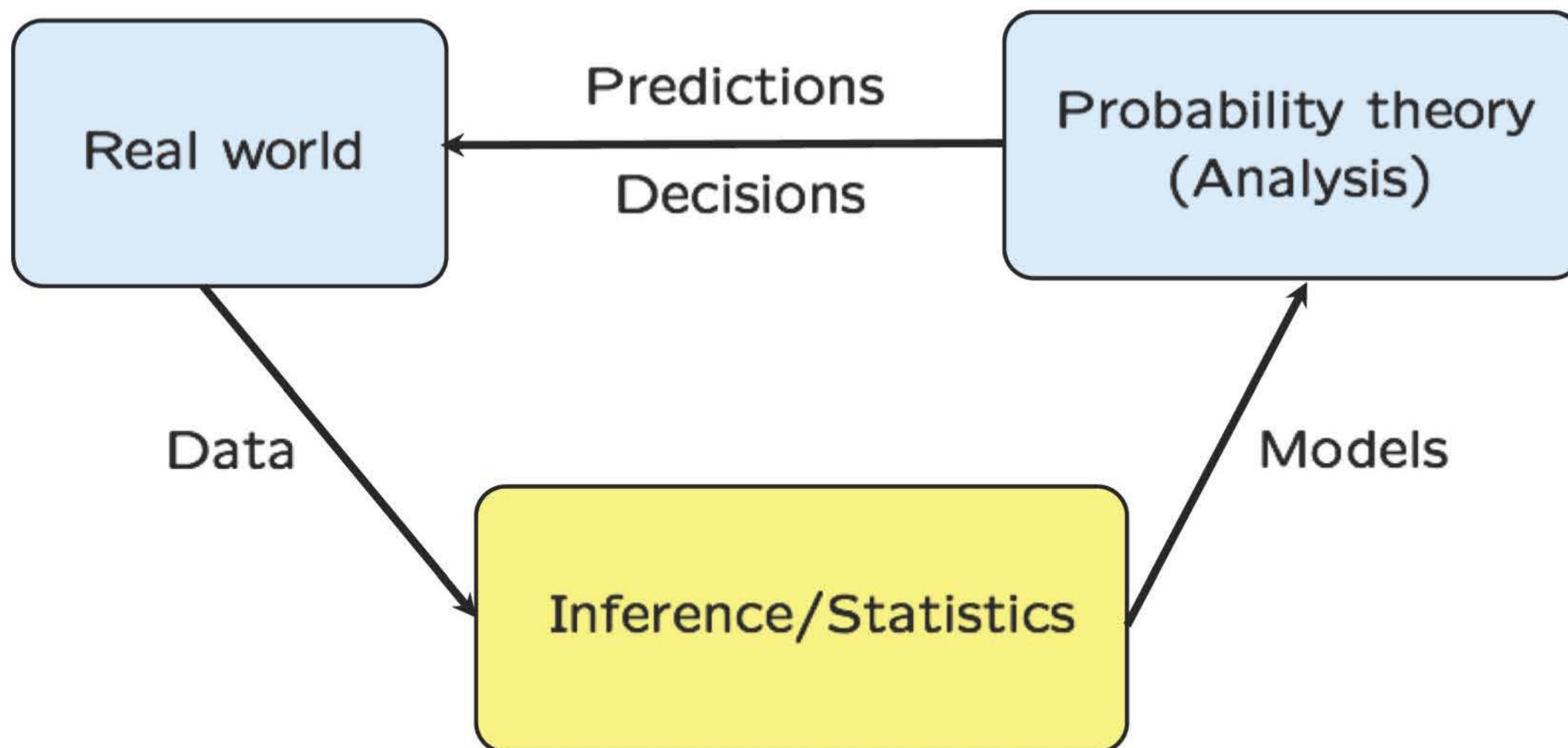


LECTURE 14: Introduction to Bayesian inference

- The big picture
 - motivation, applications
 - problem types (hypothesis testing, estimation, etc.)
- The general framework
 - Bayes' rule → posterior
(4 versions)
 - point estimates (MAP, LMS)
 - performance measures)
(prob. of error; mean squared error)
 - examples

Inference: the big picture



Inference then and now

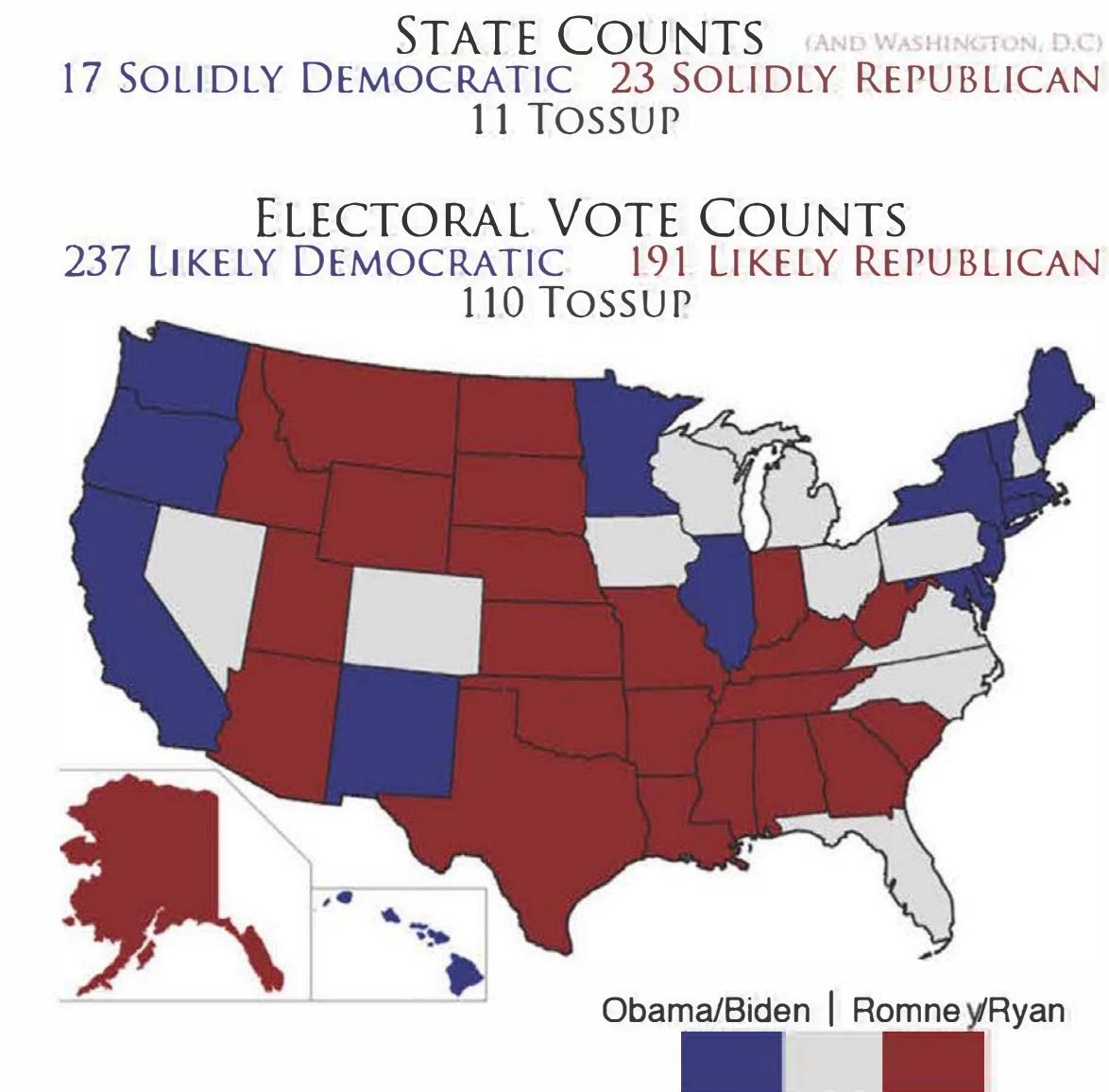
- Then:
 - 10 patients were treated: 3 died
 - 10 patients were not treated: 5 died
- Therefore ...

Now:

- Big data
- Big models
- Big computers

A sample of application domains

- Design and interpretation of experiments
 - polling



A sample of application domains

- marketing, advertising
- recommendation systems
 - Netflix competition

2	1	4	5		
5	4		?	1	3
3	5		2		
4	?		5	3	?
4	1	3		5	
2		1	?		4
1		5	5	4	
2	?	5	?	4	
3	3	1	5	2	1
3		1		2	3
4	5	1		3	
3		3	?		5
2	?	1	1		
	5	2	?	4	4
1	1	3	1	5	4
1	2	4		5	?

A sample of application domains

- Finance

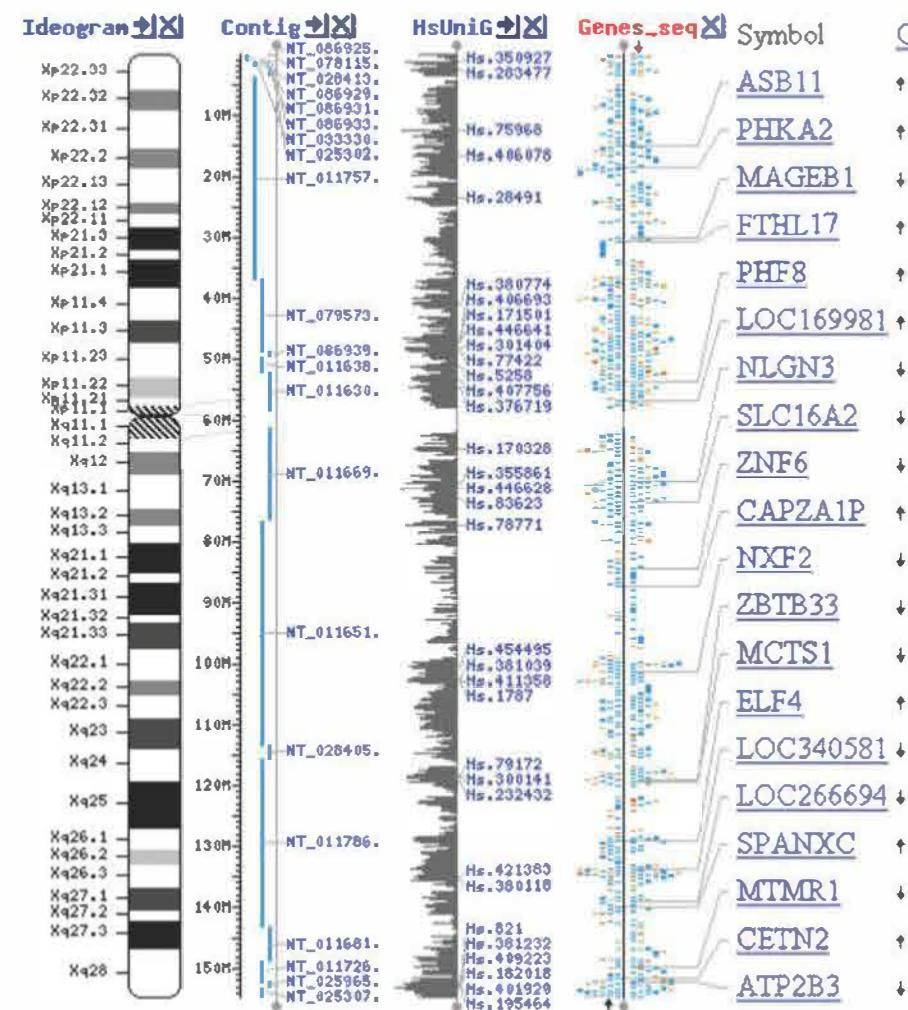


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A sample of application domains

- Life sciences

 - genomics

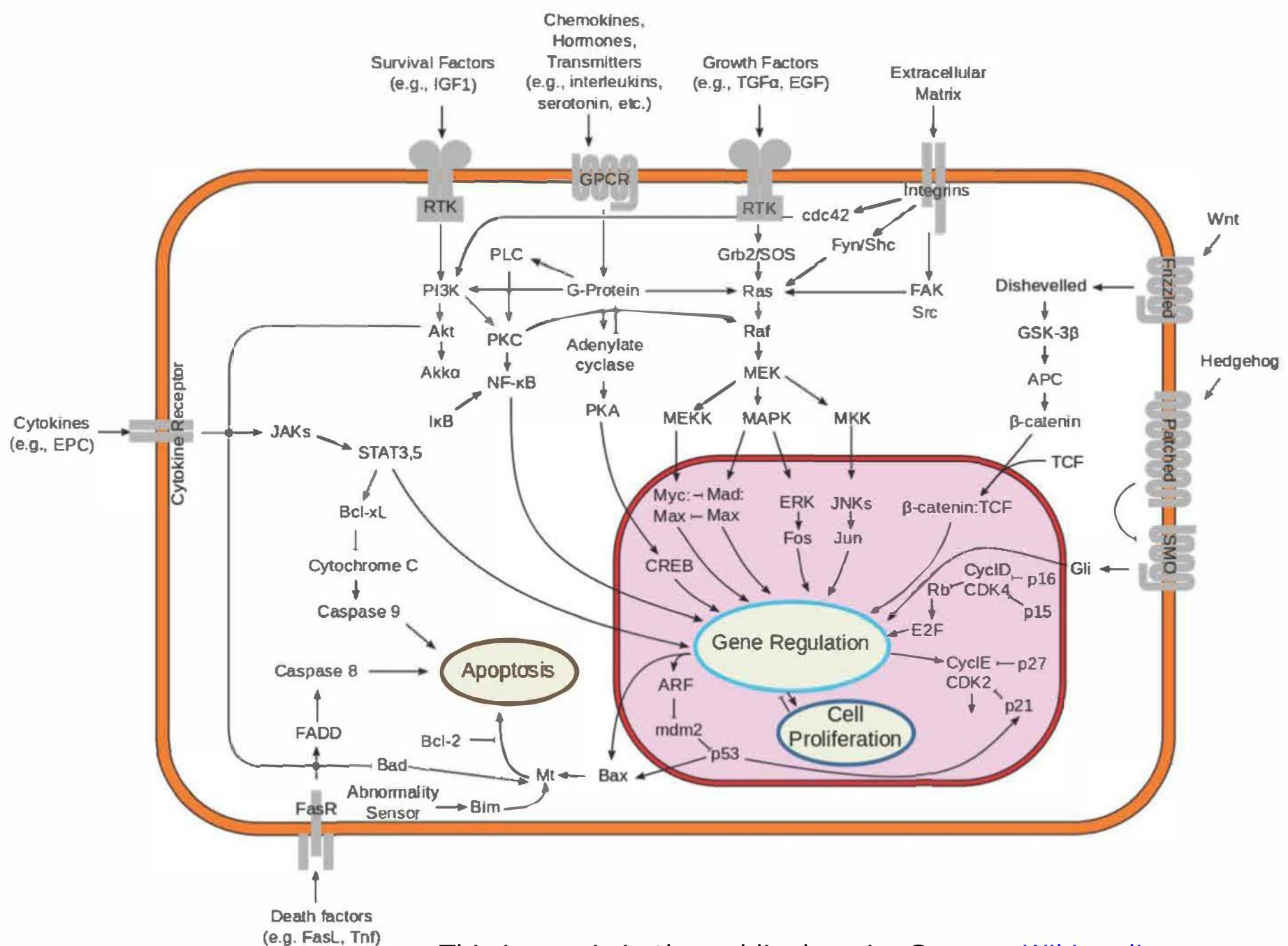


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 - neuroscience, etc., etc.

- systems biology



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A sample of application domains

- Modeling and monitoring the oceans
- Modeling and monitoring global climate
- Modeling and monitoring pollution
- Interpreting data from physics experiments
- Interpreting astronomy data

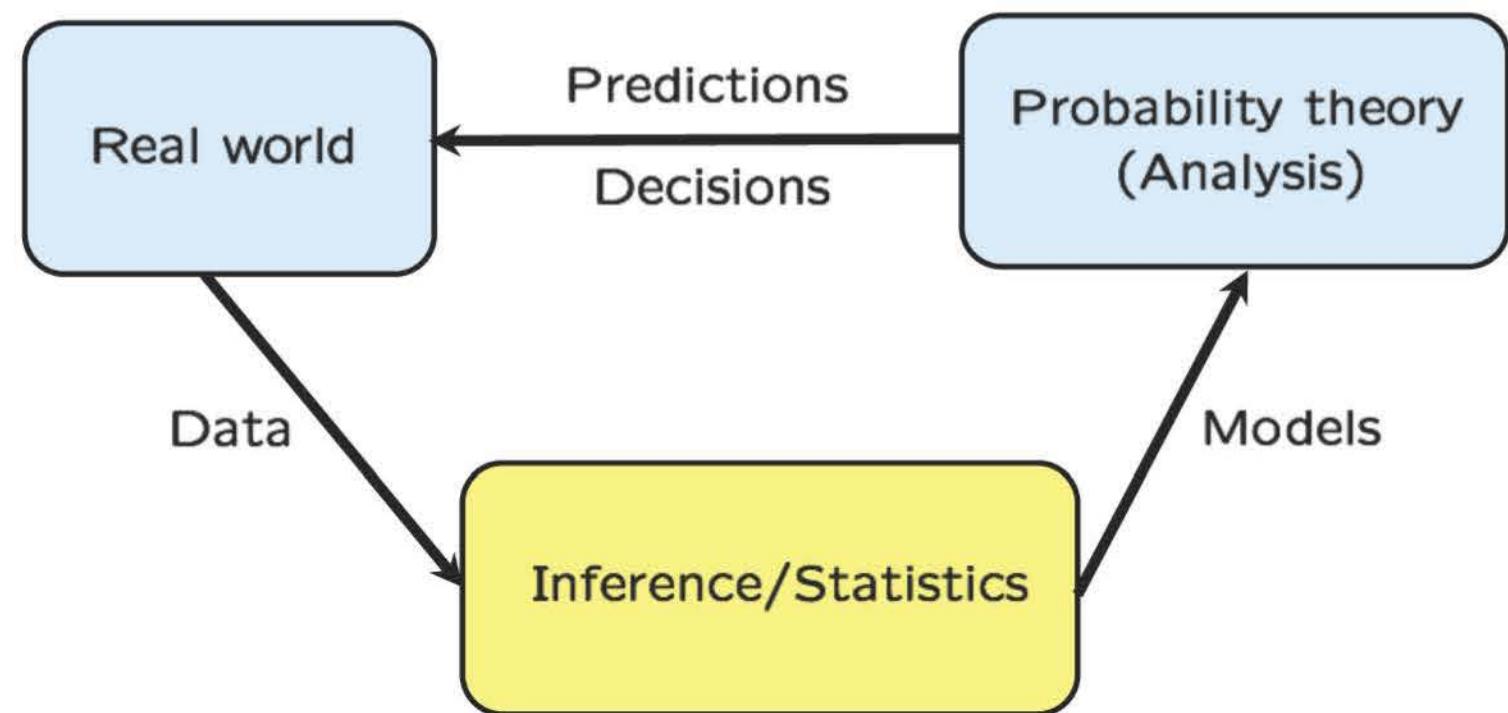
A sample of application domains

- Signal processing
 - communication systems (noisy ...)
 - speech processing and understanding
 - image processing and understanding
 - tracking of objects
 - positioning systems (e.g., GPS)
 - detection of abnormal events

Model building versus inferring unobserved variables

$$X = aS + W$$

- Model building:
 - know “signal” S , observe X
 - infer a
- Variable estimation:
 - know a , observe X
 - infer S



Hypothesis testing versus estimation

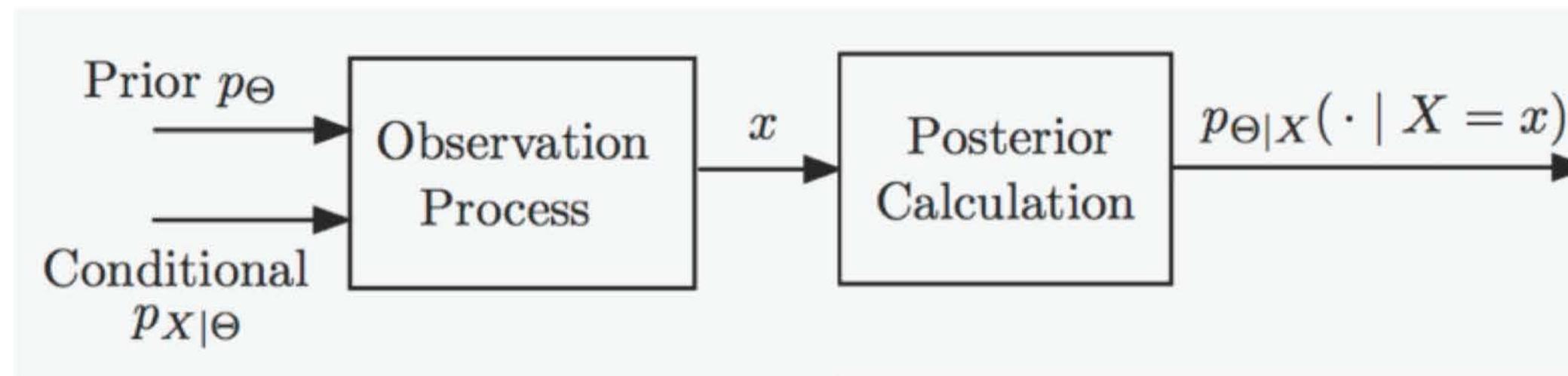
- Hypothesis testing:
 - unknown takes one of few possible values
 - aim at small probability of incorrect decision

Is it an airplane or a bird?

- Estimation:
 - numerical unknown(s)
 - aim at an estimate that is “close” to the true but unknown value

The Bayesian inference framework

- Unknown Θ
 - treated as a random variable
 - prior distribution p_Θ or f_Θ
- Observation X
 - observation model $p_{X|\Theta}$ or $f_{X|\Theta}$
- Use appropriate version of the Bayes rule to find $p_{\Theta|X}(\cdot | X = x)$ or $f_{\Theta|X}(\cdot | X = x)$
- Where does the prior come from?
 - symmetry
 - known range
 - earlier studies
 - subjective or arbitrary



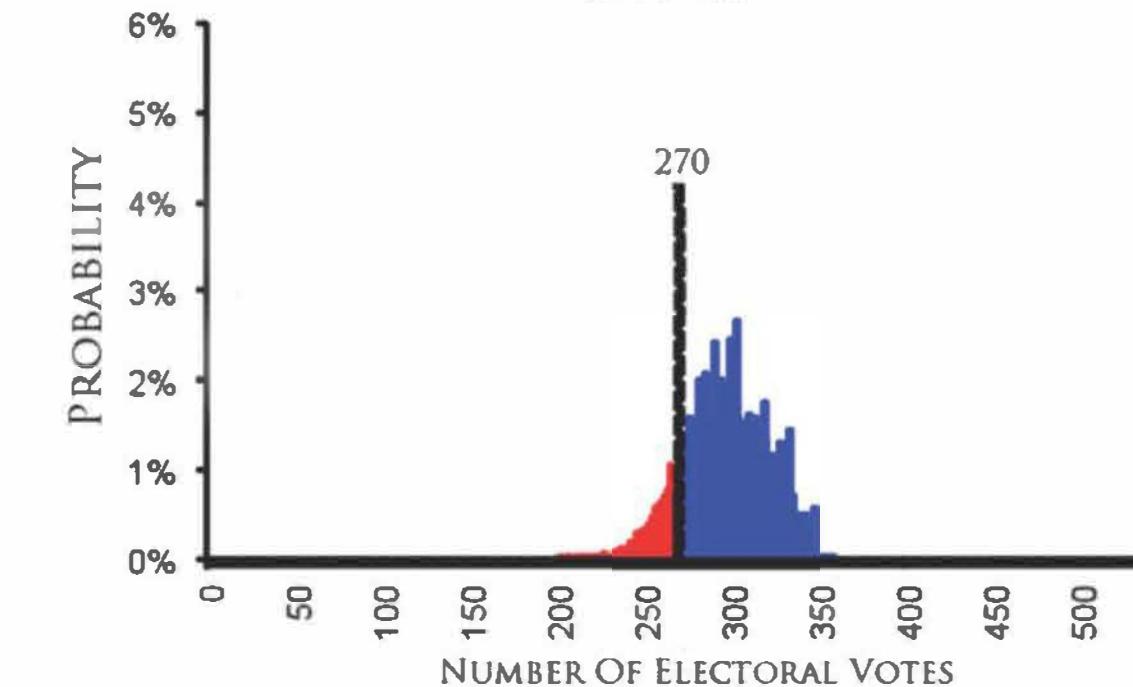
The output of Bayesian inference

The complete answer is a posterior distribution:

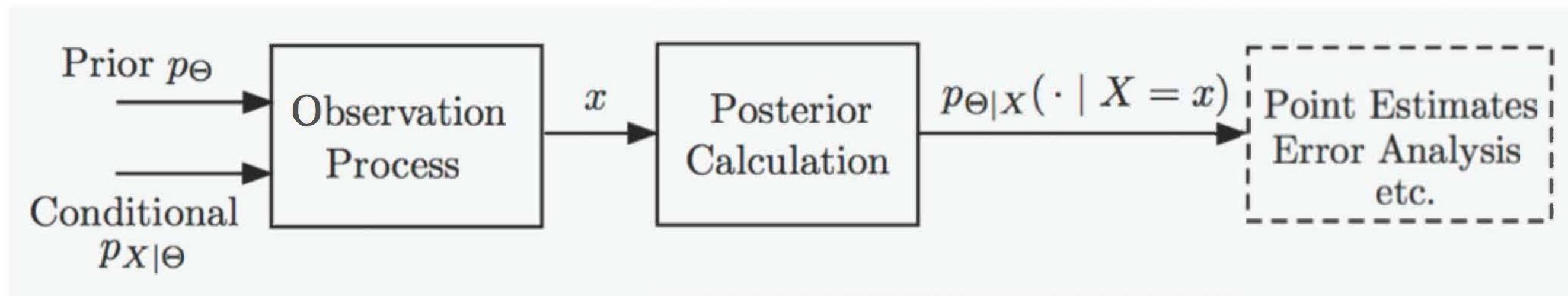
PMF $p_{\Theta|X}(\cdot | x)$ or PDF $f_{\Theta|X}(\cdot | x)$



ELECTORAL VOTE DISTRIBUTION FOR OBAMA
ROMNEY 14.62% 84.59% OBAMA
0.79% TIE



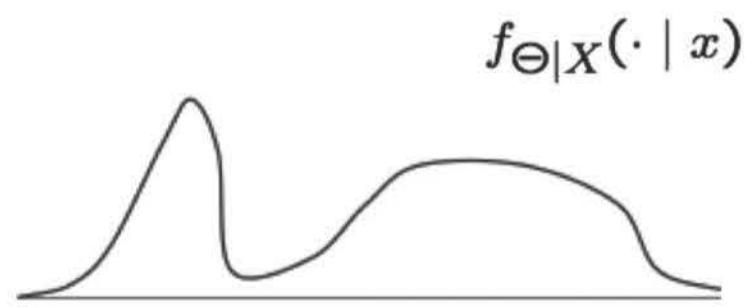
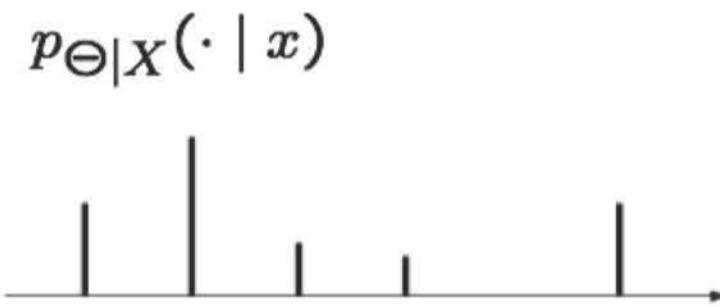
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Point estimates in Bayesian inference

The complete answer is a posterior distribution:

PMF $p_{\Theta|X}(\cdot | x)$ or PDF $f_{\Theta|X}(\cdot | x)$



estimate: $\hat{\theta} = g(x)$
(number)

estimator: $\widehat{\Theta} = g(X)$
(random variable)

- Maximum a posteriori probability (MAP):

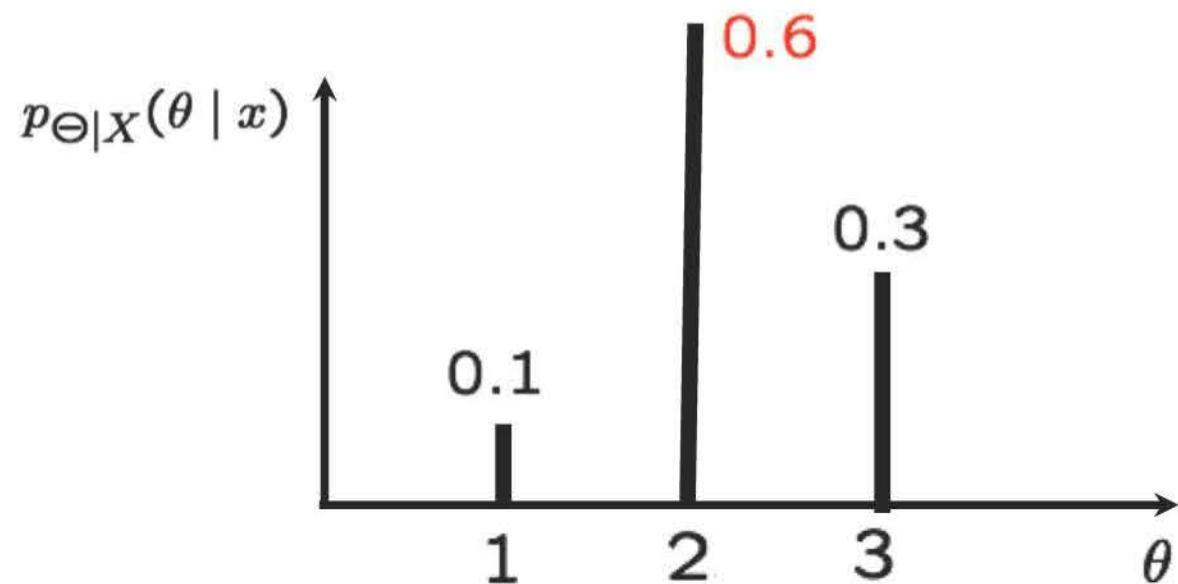
$$p_{\Theta|X}(\theta^* | x) = \max_{\theta} p_{\Theta|X}(\theta | x)$$

$$f_{\Theta|X}(\theta^* | x) = \max_{\theta} f_{\Theta|X}(\theta | x)$$

- Conditional expectation: $E[\Theta | X = x]$ (LMS: Least Mean Squares)

Discrete Θ , discrete X

- values of Θ : alternative hypotheses



$$p_{\Theta|X}(\theta | x) = \frac{p_{\Theta}(\theta) p_{X|\Theta}(x | \theta)}{p_X(x)}$$

$$p_X(x) = \sum_{\theta'} p_{\Theta}(\theta') p_{X|\Theta}(x | \theta')$$

- conditional prob of error:

$$\mathbf{P}(\hat{\theta} \neq \Theta | X = x)$$

smallest under the MAP rule

- overall probability of error:

$$\mathbf{P}(\widehat{\Theta} \neq \Theta) = \sum_x \mathbf{P}(\widehat{\Theta} \neq \Theta | X = x) p_X(x)$$

$$= \sum_{\theta} \mathbf{P}(\widehat{\Theta} \neq \Theta | \Theta = \theta) p_{\Theta}(\theta)$$

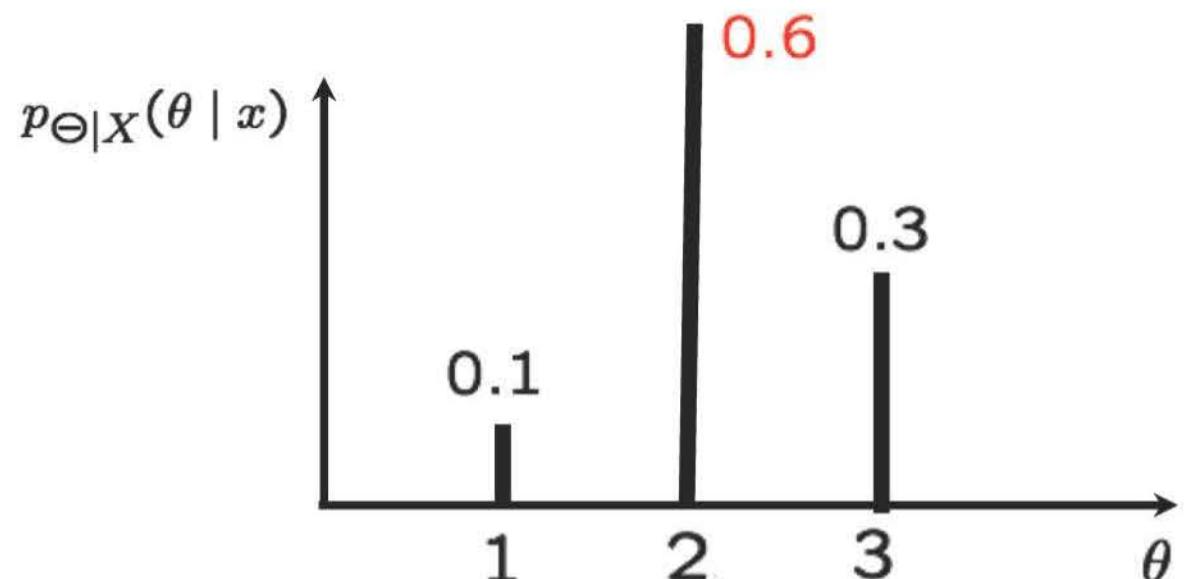
- MAP rule: $\hat{\theta} =$

Discrete Θ , continuous X

- Standard example:
 - send signal $\Theta \in \{1, 2, 3\}$
 - $X = \Theta + W$

$W \sim N(0, \sigma^2)$, indep. of Θ

$$f_{X|\Theta}(x | \theta) = f_W(x - \theta)$$



- MAP rule: $\hat{\theta} =$

$$p_{\Theta|X}(\theta | x) = \frac{p_{\Theta}(\theta) f_{X|\Theta}(x | \theta)}{f_X(x)}$$

$$f_X(x) = \sum_{\theta'} p_{\Theta}(\theta') f_{X|\Theta}(x | \theta')$$

- conditional prob of error:

$$P(\hat{\theta} \neq \Theta | X = x)$$

smallest under the MAP rule

- overall probability of error:

$$P(\hat{\Theta} \neq \Theta) = \int P(\hat{\Theta} \neq \Theta | X = x) f_X(x) dx$$

$$= \sum_{\theta} P(\hat{\Theta} \neq \theta | \Theta = \theta) p_{\Theta}(\theta)$$

Continuous Θ , continuous X

- linear normal models
estimation of a noisy signal

$$X = \Theta + W$$

Θ and W : independent normals

multi-dimensional versions (many normal parameters, many observations)

- estimating the parameter of a uniform

X : `uniform[0, Θ]`

Θ : `uniform [0, 1]`

$$f_{\Theta|X}(\theta | x) = \frac{f_{\Theta}(\theta) f_{X|\Theta}(x | \theta)}{f_X(x)}$$

$$f_X(x) = \int f_{\Theta}(\theta') f_{X|\Theta}(x | \theta') d\theta'$$

- $\widehat{\Theta} = g(X)$

- interested in:

$$\mathbf{E}[(\widehat{\Theta} - \Theta)^2 | X = x]$$

$$\mathbf{E}[(\widehat{\Theta} - \Theta)^2]$$

Inferring the unknown bias of a coin and the Beta distribution

- Standard example:
 - coin with bias Θ ; prior $f_\Theta(\cdot)$
 - fix n ; K = number of heads
- Assume $f_\Theta(\cdot)$ is uniform in $[0, 1]$

$$f_{\Theta|K}(\theta | k) = \frac{f_\Theta(\theta) p_{K|\Theta}(k | \theta)}{p_K(k)}$$

$$p_K(k) = \int f_\Theta(\theta') p_{K|\Theta}(k | \theta') d\theta'$$

$$f_{\Theta|K}(\theta | k) =$$

$$= \frac{1}{d(n, k)} \theta^k (1 - \theta)^{n-k} \quad \text{"Beta distribution, with parameters } (k+1, n-k+1)\text{"}$$

- If prior is Beta: $f_\Theta(\theta) = \frac{1}{c} \theta^\alpha (1 - \theta)^\beta$

$$f_{\Theta|K}(\theta | k) =$$

Inferring the unknown bias of a coin: point estimates

- Standard example:
 - coin with bias Θ ; prior $f_\Theta(\cdot)$
 - fix n ; K = number of heads
- Assume $f_\Theta(\cdot)$ is uniform in $[0, 1]$

$$\int_0^1 \theta^\alpha (1 - \theta)^\beta d\theta = \frac{\alpha! \beta!}{(\alpha + \beta + 1)!}$$

$$E[\Theta | K = k] =$$

$$f_{\Theta|K}(\theta | k) = \frac{1}{d(n, k)} \theta^k (1 - \theta)^{n-k}$$

- MAP estimate:

$$\hat{\theta}_{\text{MAP}} =$$

$$\widehat{\Theta}_{\text{MAP}} =$$

Summary

- Problem data: $p_{\Theta}(\cdot)$, $p_{X|\Theta}(\cdot | \cdot)$
- Given the value x of X : find, e.g., $p_{\Theta|X}(\cdot | x)$
 - using appropriate version of the Bayes rule
- Estimator $\widehat{\Theta} = g(X)$ Estimate $\widehat{\theta} = g(x)$
 - MAP: $\widehat{\theta}_{\text{MAP}} = g_{\text{MAP}}(x)$ maximizes $p_{\Theta|X}(\theta | x)$
 - LMS: $\widehat{\theta}_{\text{LMS}} = g_{\text{LMS}}(x) = \mathbb{E}[\Theta | X = x]$
- Performance evaluation of an estimator $\widehat{\Theta}$

$P(\widehat{\Theta} \neq \Theta X = x)$	$E[(\widehat{\Theta} - \Theta)^2 X = x]$
$P(\widehat{\Theta} \neq \Theta)$	$E[(\widehat{\Theta} - \Theta)^2]$

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Resource: Introduction to Probability
John Tsitsiklis and Patrick Jaillet

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