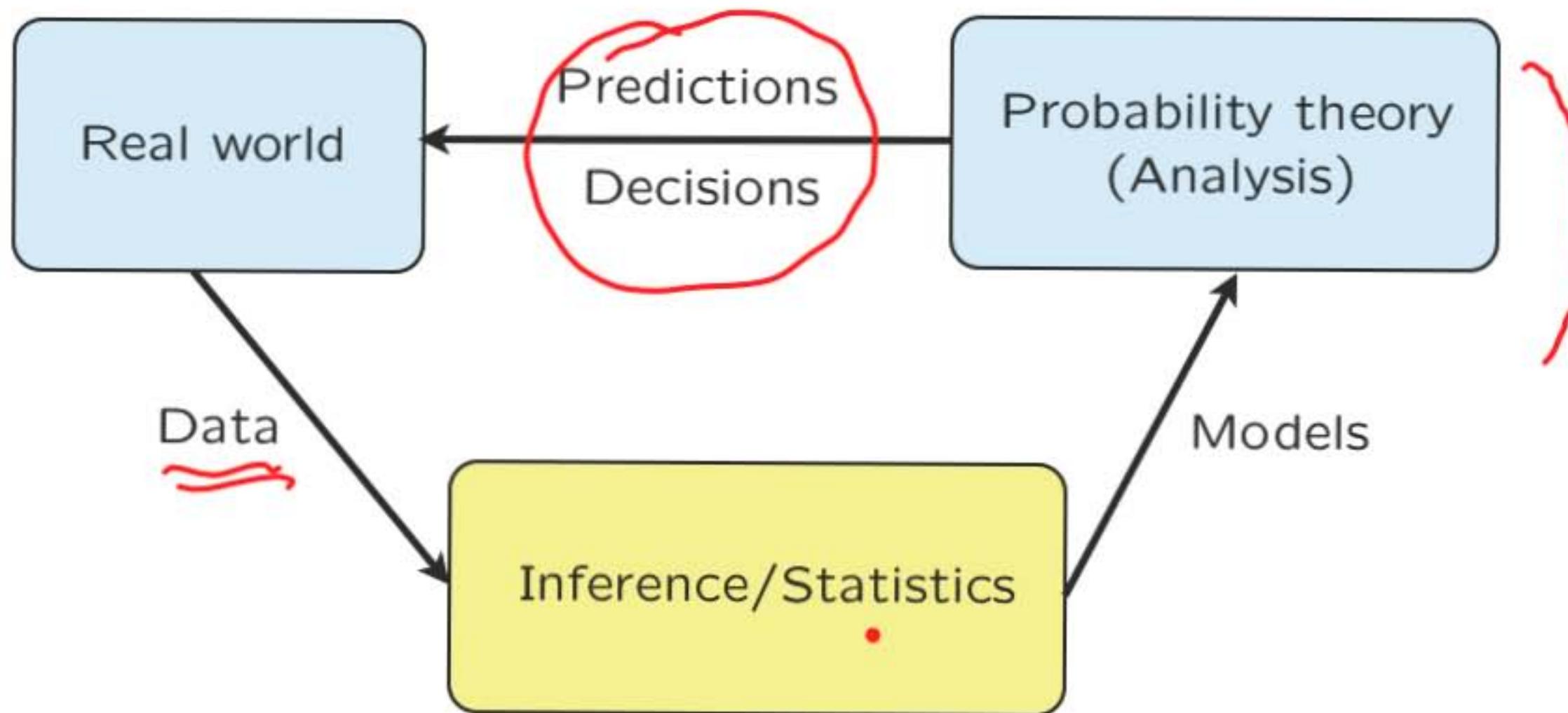


## LECTURE 14: Introduction to Bayesian inference

- The big picture
  - motivation, applications
  - problem types (hypothesis testing, estimation, etc.)
- The general framework
  - Bayes' rule → posterior  
(4 versions)
  - point estimates (MAP, LMS)
  - performance measures)  
(prob. of error; mean squared error)
  - examples

## Inference: the big picture



## Inference then and now

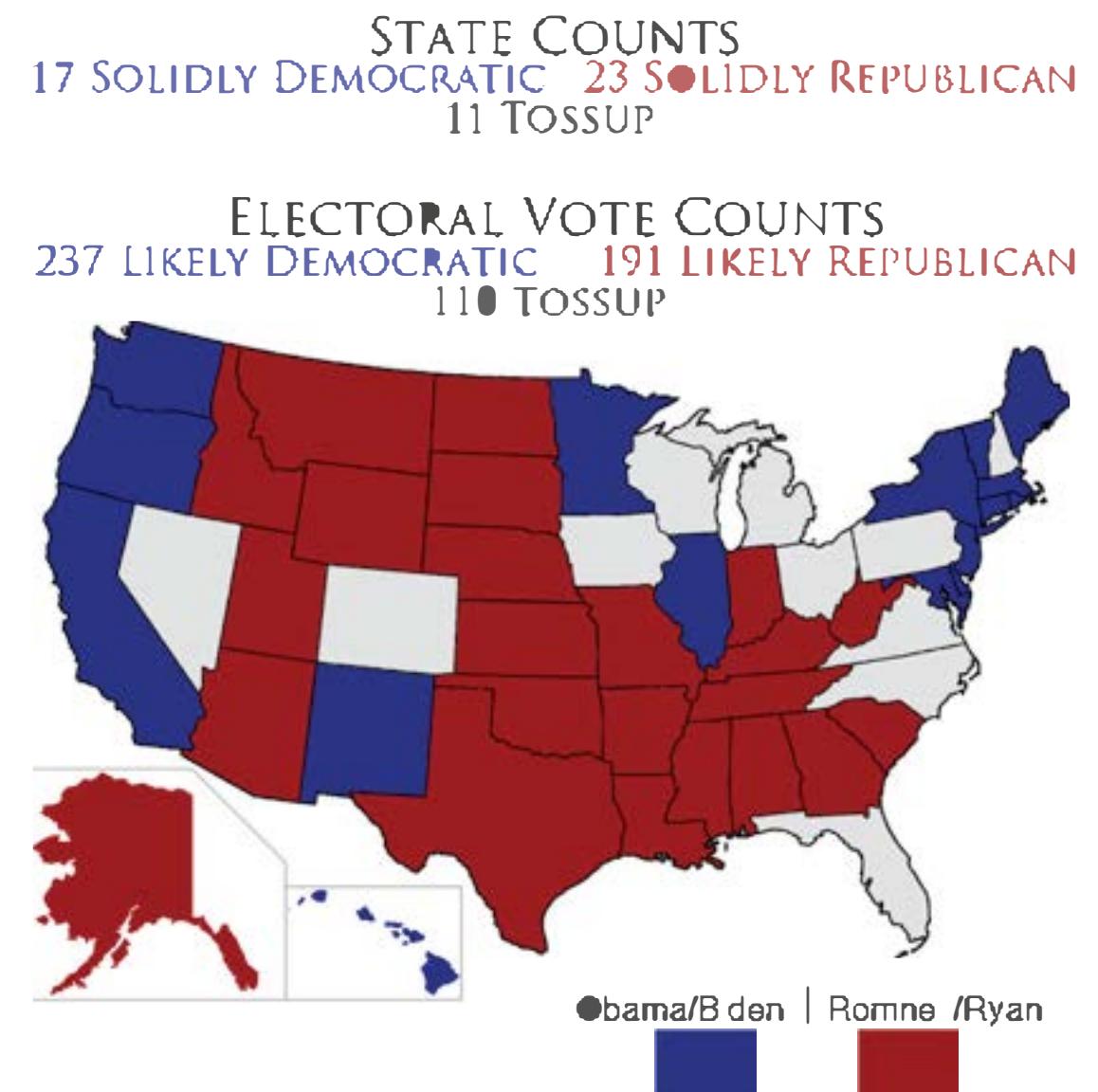
- Then:
  - 10 patients were treated: 3 died
  - 10 patients were not treated: 5 died
- Therefore ...

## Now:

- Big data
- Big models
- Big computers

# A sample of application domains

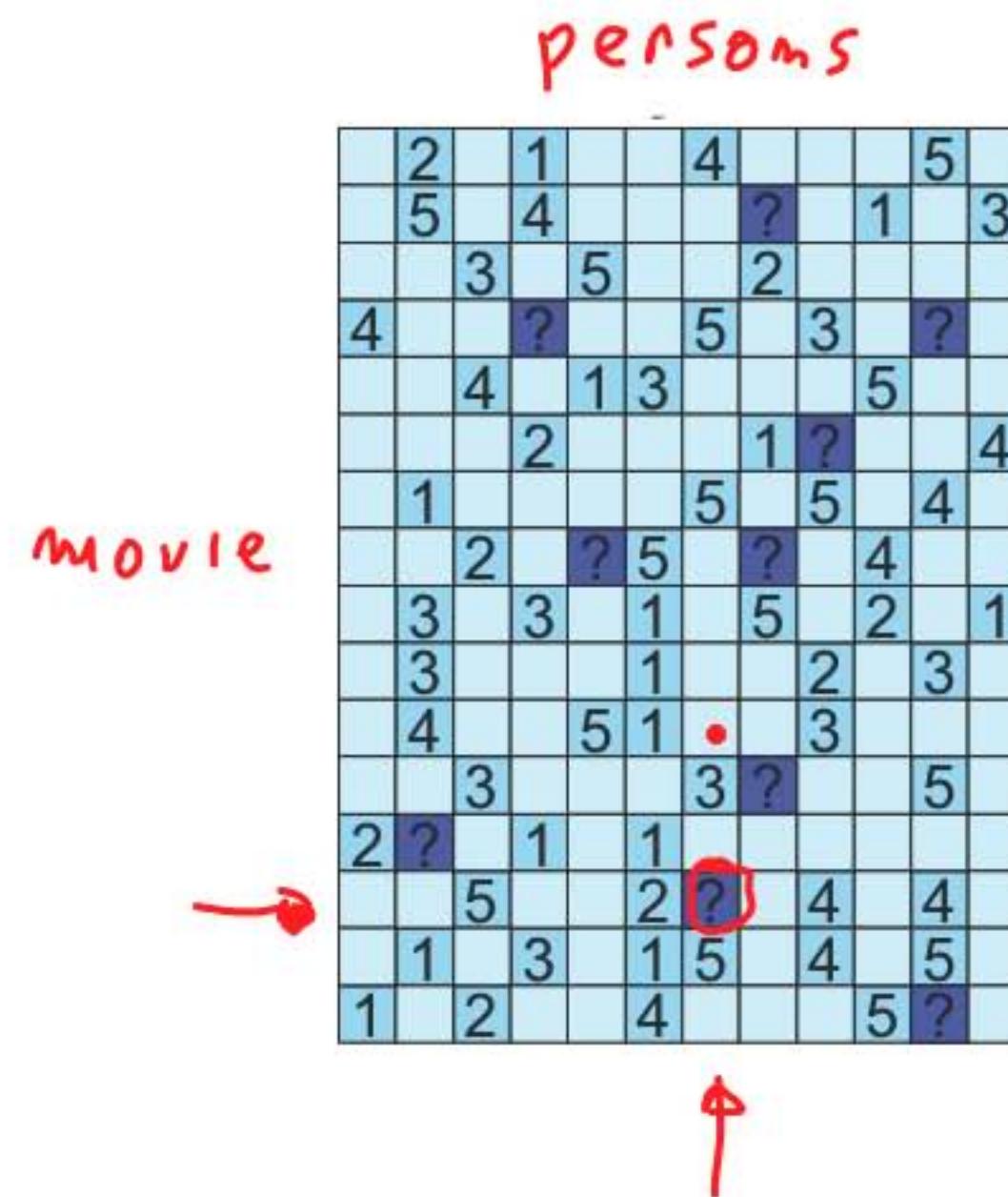
- Design and interpretation of experiments
    - polling •



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## A sample of application domains

- marketing, advertising
- recommendation systems
  - Netflix competition



## A sample of application domains

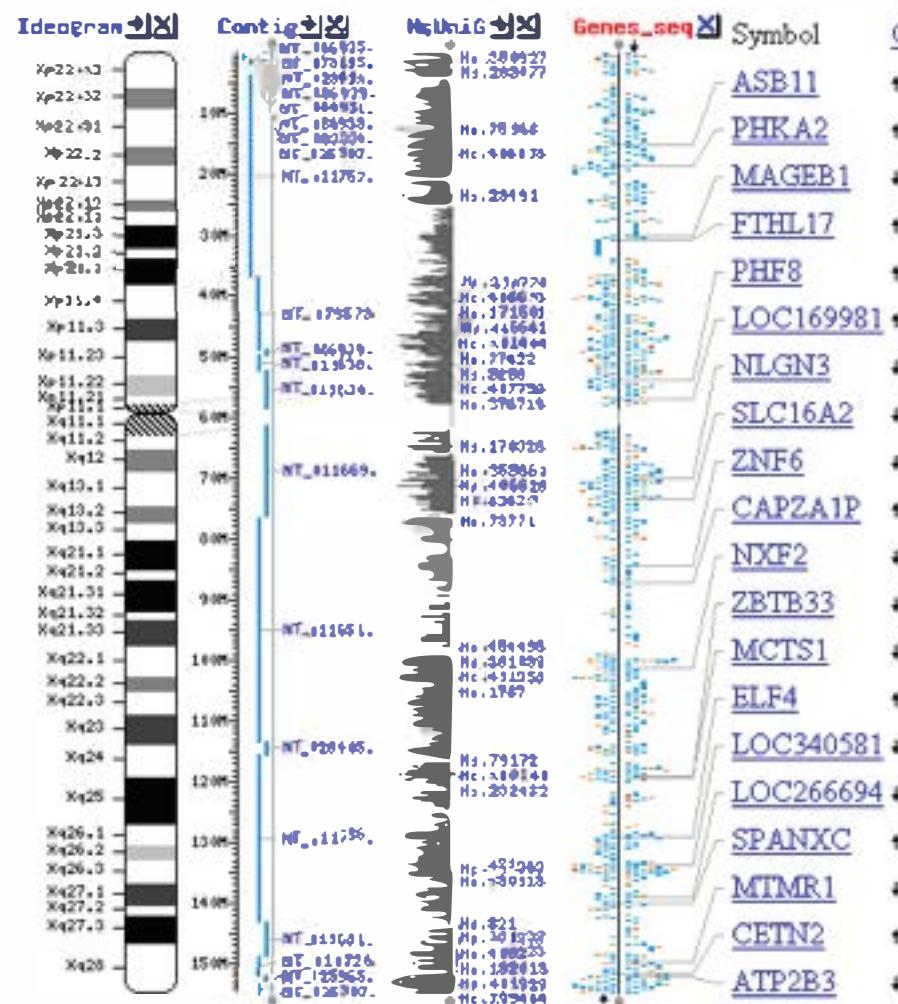
- Finance



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# A sample of application domains

- Life sciences
    - genomics

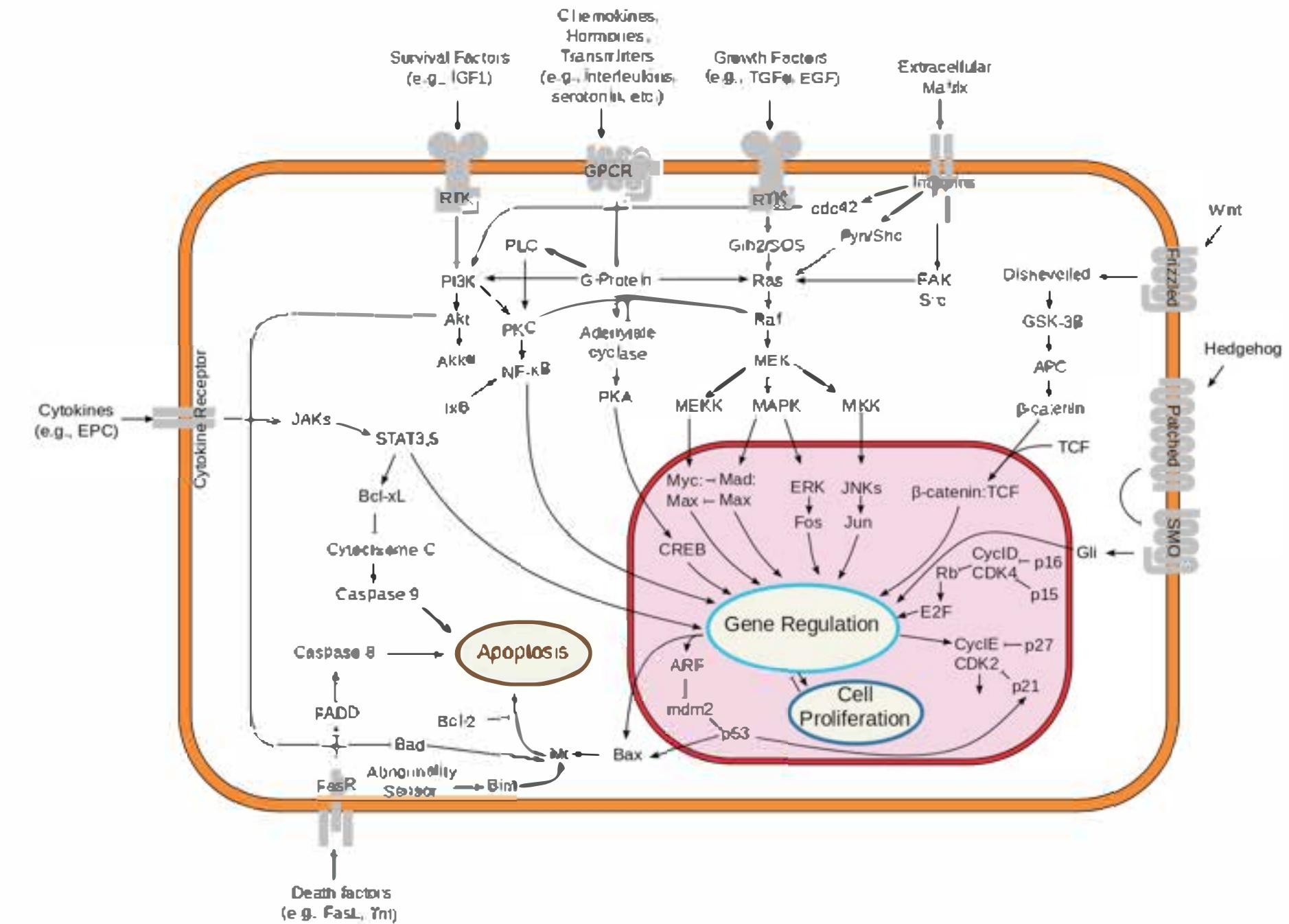


neuroscience, etc., etc.

This image is in the public domain.

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– systems biology



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## A sample of application domains

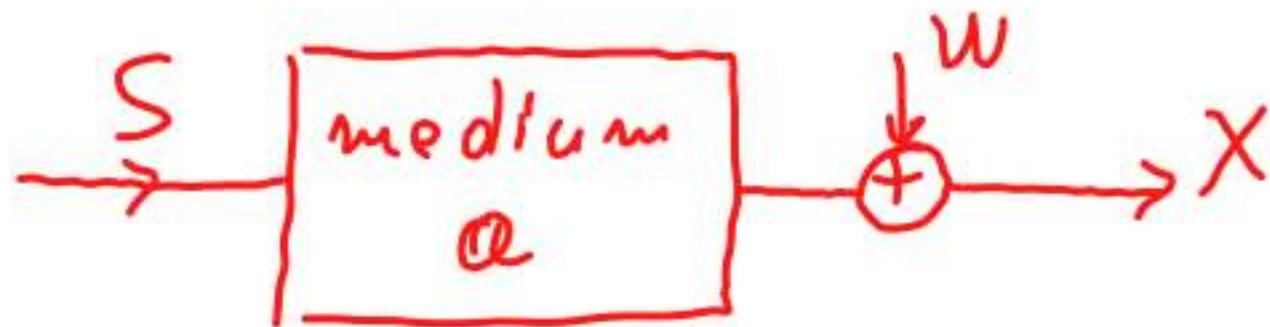
- Modeling and monitoring the oceans
- Modeling and monitoring global climate
- Modeling and monitoring pollution
- Interpreting data from physics experiments •
- Interpreting astronomy data

## A sample of application domains

- Signal processing
  - communication systems (noisy ...)
  - speech processing and understanding
  - image processing and understanding
  - tracking of objects
  - positioning systems (e.g., GPS)
  - detection of abnormal events

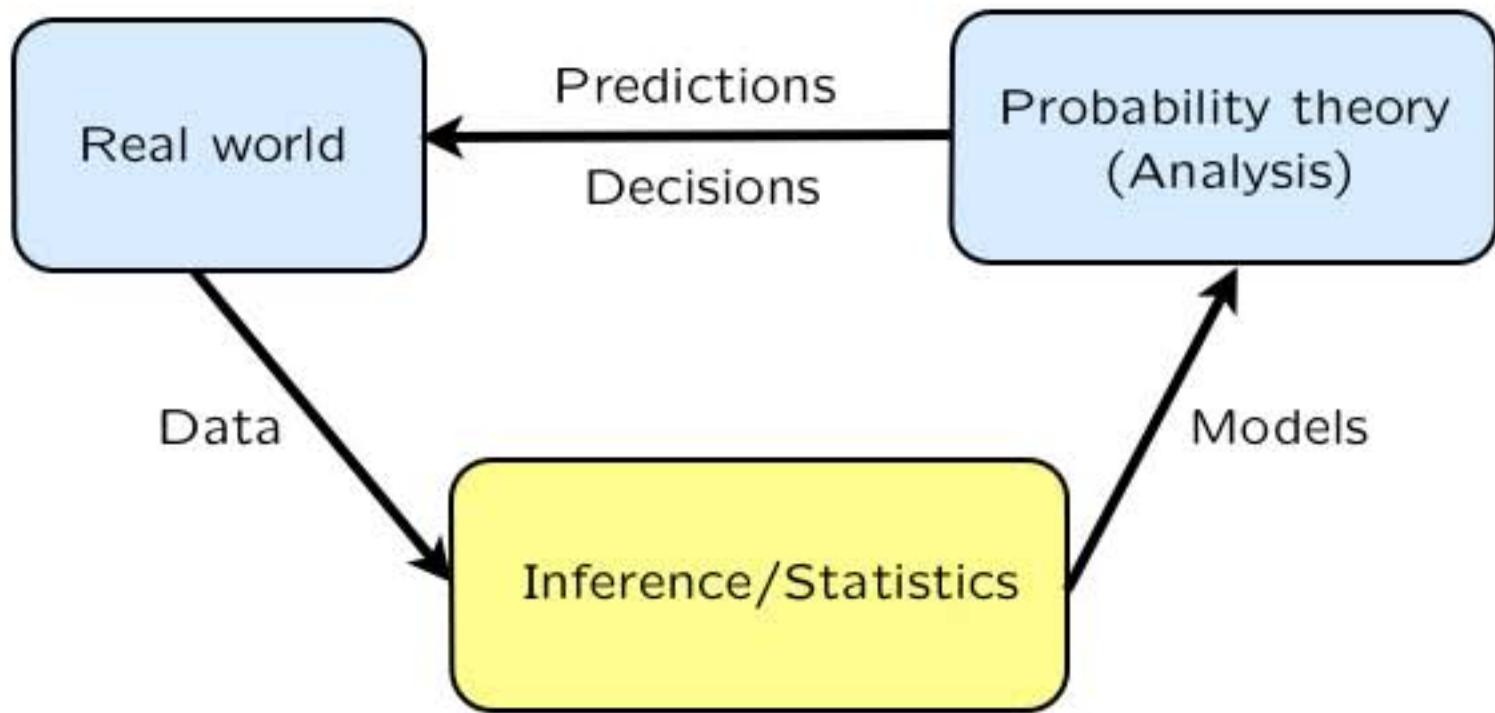
•

## Model building versus inferring unobserved variables



$$X = aS + W$$

- Model building:
  - know “signal”  $S$ , observe  $X$
  - infer  $a$
- Variable estimation:
  - know  $a$ , observe  $X$
  - infer  $S$



## Hypothesis testing versus estimation

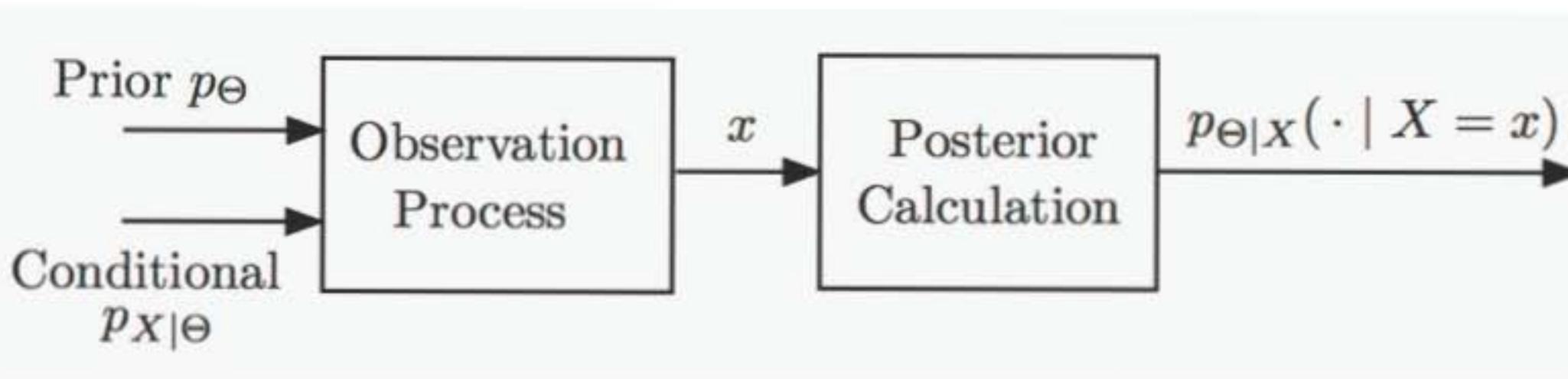
- Hypothesis testing:
  - unknown takes one of few possible values
  - aim at small probability of incorrect decision

Is it an airplane or a bird?

- Estimation:
  - numerical unknown(s)
  - aim at an estimate that is “close” to the true but unknown value

## The Bayesian inference framework

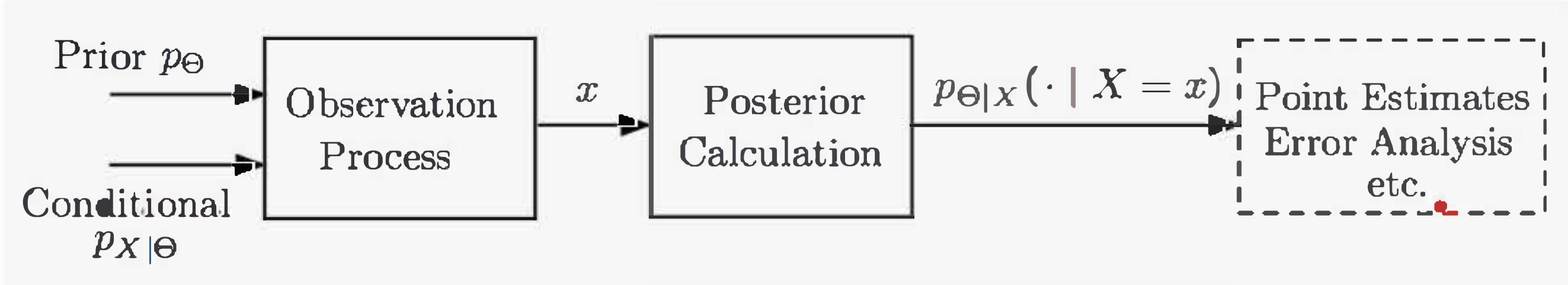
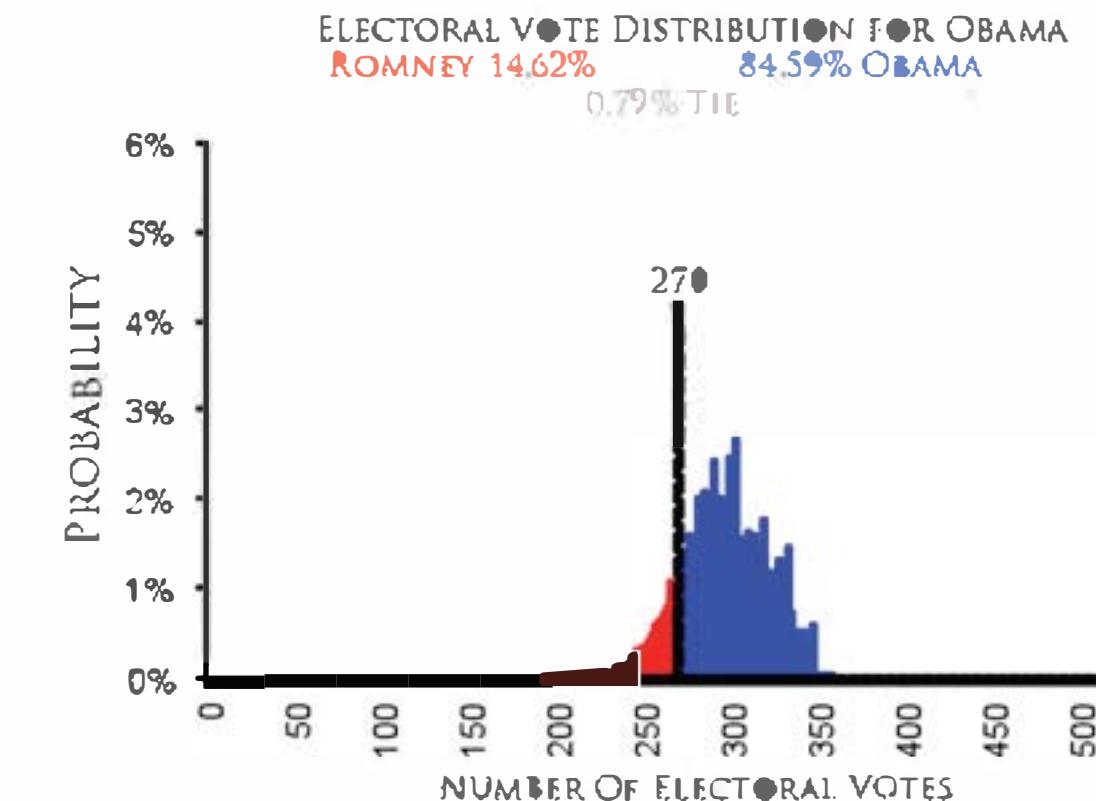
- Unknown  $\Theta$ 
  - treated as a random variable
  - prior distribution  $p_\Theta$  or  $f_\Theta$
- Observation  $X$ 
  - observation model  $p_{X|\Theta}$  or  $f_{X|\Theta}$
- Use appropriate version of the Bayes rule to find  $p_{\Theta|X}(\cdot | X = x)$  or  $f_{\Theta|X}(\cdot | X = x)$
- Where does the prior come from?
  - symmetry
  - known range
  - earlier studies
  - subjective or arbitrary



# The output of Bayesian inference

The complete answer is a posterior distribution:

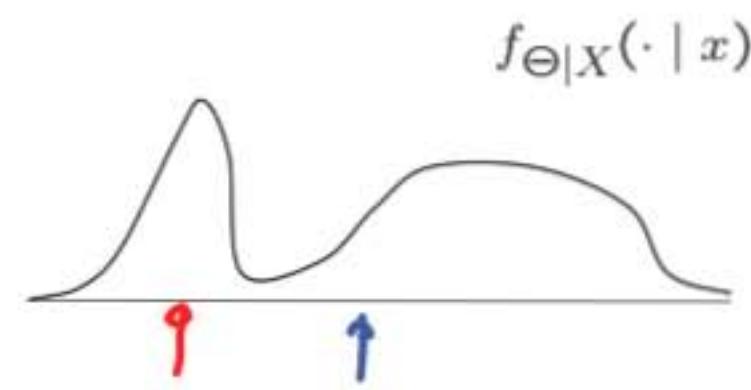
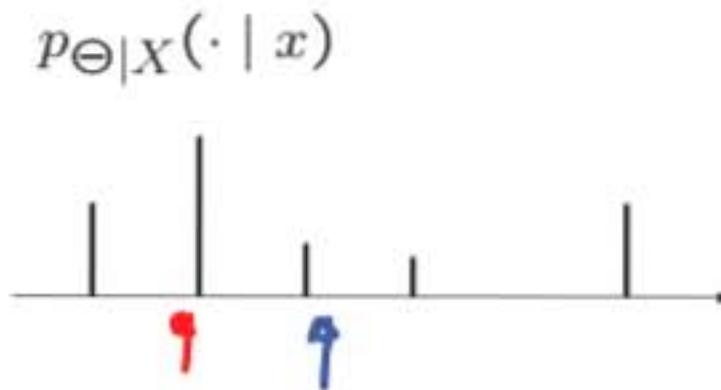
PMF  $p_{\Theta|X}(\cdot | x)$  or PDF  $f_{\Theta|X}(\cdot | x)$



## Point estimates in Bayesian inference

The complete answer is a posterior distribution:

PMF  $p_{\Theta|X}(\cdot | x)$  or PDF  $f_{\Theta|X}(\cdot | x)$



**estimate:**  $\hat{\theta} = g(x)$   
(number)

**estimator:**  $\widehat{\Theta} = g(X)$   
(random variable)

- Maximum a posteriori probability (MAP):

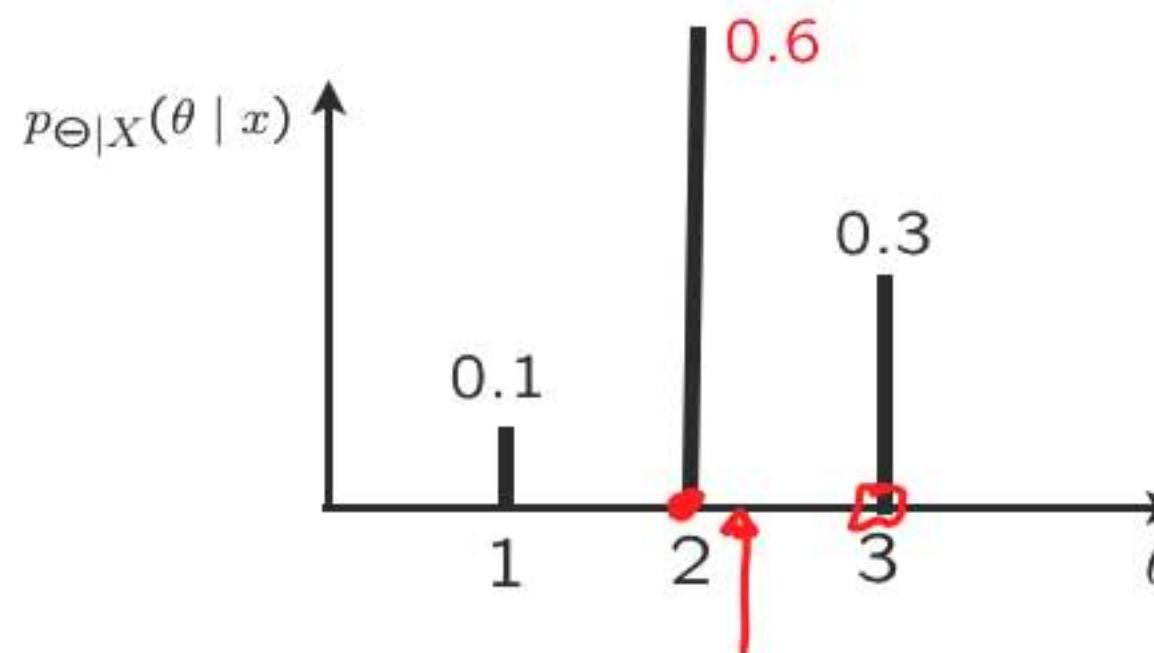
$$p_{\Theta|X}(\theta^* | x) = \max_{\theta} p_{\Theta|X}(\theta | x),$$

$$f_{\Theta|X}(\theta^* | x) = \max_{\theta} f_{\Theta|X}(\theta | x).$$

- Conditional expectation:  $E[\Theta | X = x]$  (LMS: Least Mean Squares)

## Discrete $\Theta$ , discrete $X$

- values of  $\Theta$ : alternative hypotheses



- MAP rule:  $\hat{\theta} = 2$

$$LMS: \hat{\theta} = E[\Theta | X=x] = 2.2$$

$$p_{\Theta|X}(\theta | x) = \frac{p_{\Theta}(\theta) p_{X|\Theta}(x | \theta)}{p_X(x)}$$

$$p_X(x) = \sum_{\theta'} p_{\Theta}(\theta') p_{X|\Theta}(x | \theta')$$

- conditional prob of error:

$$P(\hat{\theta} \neq \Theta | X=x) = 0.4$$

smallest under the MAP rule

- overall probability of error:

$$P(\hat{\Theta} \neq \Theta) = \sum_x P(\hat{\Theta} \neq \Theta | X=x) p_X(x)$$

$$= \sum_{\theta} P(\hat{\Theta} \neq \Theta | \Theta = \theta) p_{\Theta}(\theta)$$

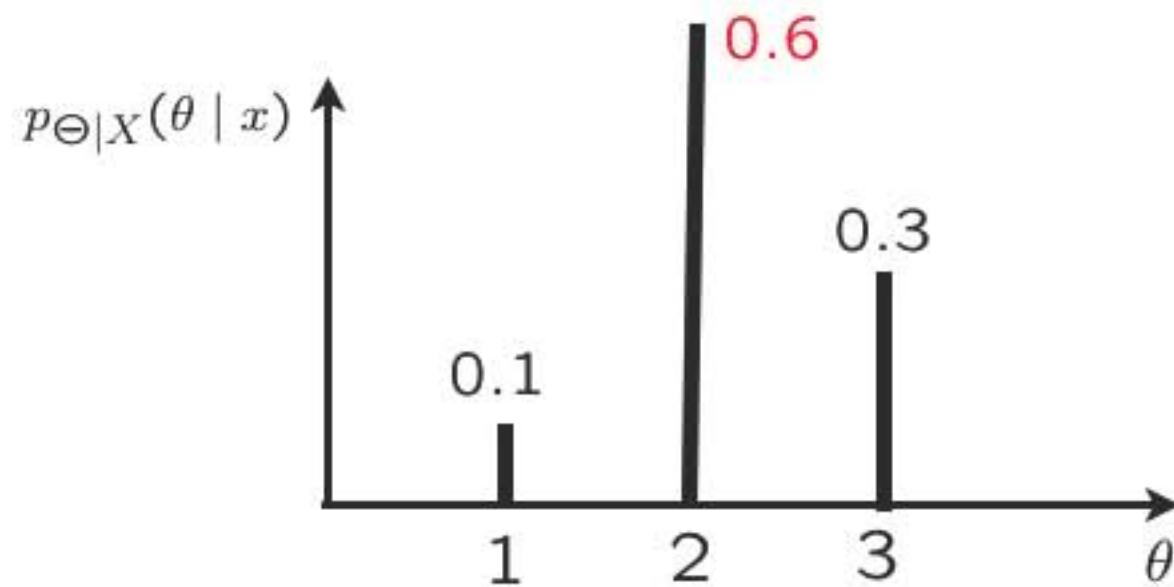
## Discrete $\Theta$ , continuous $X$

- Standard example:
  - send signal  $\Theta \in \{1, 2, 3\}$

$$X = \Theta + W$$

$W \sim N(0, \sigma^2)$ , indep. of  $\Theta$

$$f_{X|\Theta}(x | \theta) = f_W(x - \theta)$$



- MAP rule:  $\hat{\theta} = 2$

$$p_{\Theta|X}(\theta | x) = \frac{p_{\Theta}(\theta) f_{X|\Theta}(x | \theta)}{f_X(x)}$$

$$f_X(x) = \sum_{\theta'} p_{\Theta}(\theta') f_{X|\Theta}(x | \theta')$$

- conditional prob of error:

$$\mathbf{P}(\hat{\theta} \neq \Theta | X = x)$$

- smallest under the MAP rule
- overall probability of error:

$$\begin{aligned} \mathbf{P}(\hat{\Theta} \neq \Theta) &= \int \underbrace{\mathbf{P}(\hat{\Theta} \neq \Theta | X = x)}_{=} f_X(x) dx \\ &= \sum_{\theta} \mathbf{P}(\hat{\Theta} \neq \theta | \Theta = \theta) p_{\Theta}(\theta) \end{aligned}$$

## Continuous $\Theta$ , continuous $X$

- linear normal models  
estimation of a noisy signal

$$X = \Theta + W$$

$\Theta$  and  $W$ : independent normals

multi-dimensional versions (many normal parameters, many observations)

- estimating the parameter of a uniform

$X$ : `uniform[0, Θ]`

$\Theta$ : `uniform [0, 1]`

$$f_{\Theta|X}(\theta | x) = \frac{f_{\Theta}(\theta) f_{X|\Theta}(x | \theta)}{f_X(x)}$$

$$f_X(x) = \int f_{\Theta}(\theta') f_{X|\Theta}(x | \theta') d\theta'$$

- $\widehat{\Theta} = g(X)$   $MAP$   
 $LMS$
- interested in:

$$\left\{ \begin{array}{l} E[(\widehat{\Theta} - \Theta)^2 | X = x] \\ E[(\widehat{\Theta} - \Theta)^2] \end{array} \right.$$

## Inferring the unknown bias of a coin and the Beta distribution

- Standard example:
  - coin with bias  $\Theta$ ; prior  $f_\Theta(\cdot)$
  - fix  $n$ ;  $K$  = number of heads
- Assume  $f_\Theta(\cdot)$  is uniform in  $[0, 1]$

$$f_{\Theta|K}(\theta | k) = \frac{f_\Theta(\theta) p_{K|\Theta}(k | \theta)}{p_K(k)}$$

$$p_K(k) = \int f_\Theta(\theta') p_{K|\Theta}(k | \theta') d\theta'$$

$$f_{\Theta|K}(\theta | k) = \frac{1 \cdot \binom{n}{k} \theta^k (1-\theta)^{n-k}}{p_K(k)}$$

$$= \frac{1}{d(n, k)} \theta^k (1-\theta)^{n-k}$$

“Beta distribution, with parameters  $(k+1, n-k+1)$ ”

$$\theta \in [0, 1]$$

- If prior is Beta:  $f_\Theta(\theta) = \frac{1}{c} \theta^\alpha (1-\theta)^\beta$   $\alpha, \beta > 0$

$$f_{\Theta|K}(\theta | k) = \frac{\frac{1}{c} \theta^\alpha (1-\theta)^\beta \binom{n}{k} \theta^k (1-\theta)^{n-k}}{p_K(k)} = d \theta^{\alpha+k} (1-\theta)^{\beta+n-k}$$

## Inferring the unknown bias of a coin: point estimates

- Standard example:
  - coin with bias  $\Theta$ ; prior  $f_\Theta(\cdot)$
  - fix  $n$ ;  $K$  = number of heads
- Assume  $f_\Theta(\cdot)$  is uniform in  $[0, 1]$

$$f_{\Theta|K}(\theta | k) = \frac{1}{d(n, k)} \underline{\theta^k (1-\theta)^{n-k}}$$

- MAP estimate:

$$\hat{\theta}_{\text{MAP}} = \boxed{k/n}$$

$$\max_{\theta} [k \log \theta + (n-k) \log(1-\theta)]$$

$$\frac{k}{\theta} + (n-k)/(1-\theta) = 0$$

$$\hat{\Theta}_{\text{MAP}} = \boxed{k/n}$$

$$\int_0^1 \theta^\alpha (1-\theta)^\beta d\theta = \frac{\alpha! \beta!}{(\alpha+\beta+1)!}$$

$\alpha \geq 0$   
 $\beta \geq 0$

$$\begin{aligned} E[\Theta | K = k] &= \int_0^1 \theta f_{\Theta|K}(\theta | k) d\theta \\ &= \frac{1}{d(n, k)} \int_0^1 \theta^{k+1} (1-\theta)^{n-k} d\theta \\ &= \frac{1}{\cancel{k!} \cancel{(n-k)!} (n+1)!} \cdot \frac{(k+1)! (n-k)!}{(n+2)!} \\ &= \boxed{\frac{k+1}{n+2}} \approx \frac{k}{n} \quad (\text{if } n \text{ large}) \end{aligned}$$

## Summary

- Problem data:  $p_{\Theta}(\cdot)$ ,  $p_{X|\Theta}(\cdot | \cdot)$
- Given the value  $x$  of  $X$ : find, e.g.,  $p_{\Theta|X}(\cdot | x)$ 
  - using appropriate version of the Bayes rule (4 choices)
- Estimator  $\widehat{\Theta} = g(X)$       Estimate  $\widehat{\theta} = g(x)$ 
  - MAP:  $\widehat{\theta}_{\text{MAP}} = g_{\text{MAP}}(x)$  maximizes  $p_{\Theta|X}(\theta | x)$
  - LMS:  $\widehat{\theta}_{\text{LMS}} = g_{\text{LMS}}(x) = \mathbb{E}[\Theta | X = x]$

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Resource: Introduction to Probability  
John Tsitsiklis and Patrick Jaillet

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